



# The ability of the Global Leadership Initiative on Malnutrition (GLIM) to diagnose protein–energy malnutrition in patients requiring vascular surgery: a validation study

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## Abstract

Identifying nutritional deficits and implementing appropriate interventions in patients requiring vascular surgery is challenging due to the paucity of appropriate screening and assessment tools in this group. This retrospective study aimed to determine the validity of the Global Leadership Initiative on Malnutrition (GLIM) in identifying protein–energy malnutrition (PEM) in inpatients admitted to a vascular surgery unit, using the PG-SGA as the comparator. Diagnostic accuracy and consistency were determined between the GLIM and the Patient-Generated Subjective Global Assessment (PG-SGA) global rating. The GLIM determination was made retrospectively using the relevant parameters collected at baseline in the original study. Two hundred and twenty-four (70.1 % male) participants were included. The prevalence of PEM was 28.6 % on GLIM and 17 % via the PG-SGA. Compared with the PG-SGA, the GLIM achieved sensitivity of 73.7 % and specificity of 80.6 %; however positive predictive value was 43.7 % indicating that the GLIM over-diagnosed malnutrition compared with the PG-SGA. Kappa reached 0.427 indicating moderate diagnostic consistency. Due to the absence of an ideal instrument and the complexity of malnutrition often seen in this group which extends beyond PEM to significant micronutrient deficiencies, further work is required to determine the most appropriate instrument in this patient group, and how micronutrient status can also be included in the overall assessment given the critical role of micronutrients in this group.

**Key words:** Malnutrition; Vascular surgery; Global leadership initiative on malnutrition; GLIM; Validation

Patients admitted to vascular surgery units are a nutritionally vulnerable group with rates of malnutrition as high as 60–90 % cited in the literature<sup>(1–4)</sup>. Poor nutritional health has significant consequences such as higher rates of infection<sup>(5)</sup>, longer hospital length of stay<sup>(6,7)</sup> and more proximal amputations in those with diabetic foot infections<sup>(7)</sup>.

Identification and nutritional management of malnutrition in patients admitted to vascular surgery units is paramount to maximise nutritional health and clinical outcomes. To date, there has been limited research examining methods to identify and diagnose malnutrition in this patient group. We have previously reported that four commonly used malnutrition screening tools (Malnutrition Universal Screening Tool, Malnutrition Screening Tool, Nutrition Risk Screen-2002 and the Mini Nutritional Assessment – Short Form) were ineffective in identifying risk of malnutrition amongst vascular surgery patients<sup>(8)</sup>. However

very little evidence is available regarding appropriate nutrition assessment tools.

In 2019, the Global Leadership Initiative on Malnutrition (GLIM) was proposed as a diagnostic framework for diagnosing protein–energy malnutrition (PEM). The purpose of GLIM was to build a global consensus regarding the criteria required for diagnosing PEM in a clinical setting. Empirical consensus was reached that the first step of GLIM is using a validated screening tool to identify patients at risk of malnutrition. The next step, the diagnosis of PEM, is derived from the presence of one or more of three phenotypic criteria (non-intentional weight loss, low BMI and reduced muscle mass) and one or more of two aetiological criteria (reduced food intake and inflammation/disease burden)<sup>(9)</sup>. Work has commenced on validating the GLIM framework with a recent study in adult inpatients examining the performance of GLIM using Subjective Global Assessment

**Abbreviations:** GLIM, Global Leadership Initiative on Malnutrition; PG-SGA, Patient-Generated Subjective Global Assessment; PEM, protein–energy malnutrition; SGA, Subjective Global Assessment; Sn, sensitivity; Sp, specificity.

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(SGA) as the comparator<sup>(10)</sup>. Results showed sensitivity (Sn) and specificity (Sp) values of 61.3% and 89.8%, respectively, when GLIM malnourished was compared with malnourished (SGA B and C combined) on the SGA. Similar work has been undertaken in a range of clinical specialties included geriatric rehabilitation<sup>(11)</sup>, ambulatory cancer care<sup>(12)</sup> and intensive care patients<sup>(13,14)</sup> with Sn values of 56.7%–100% and Sp of 55.3%–98.1%. Studies examining the GLIM vary in the interpretation of the GLIM process, including how the criteria are interpreted as well as the inclusion or exclusion of the initial screening.

A nutrition assessment tool commonly used in the clinical setting is the Patient-Generated Subjective Global Assessment (PG-SGA)<sup>(15)</sup>. The PG-SGA incorporates a range of parameters to determine whether a patient is well nourished (PG-SGA-A), suspected or moderately malnourished (PG-SGA-B) or severely malnourished (PG-SGA-C) and has been used as the gold standard in recent studies exploring the validity of the GLIM<sup>(16,17)</sup>. In these studies, agreement between the GLIM and PG-SGA was low<sup>(17)</sup> to fair<sup>(16)</sup> with Sn of 43% and 51% and Sp of 79% and 98% and kappa ( $\kappa$ ) of 0.22 and 0.37. Rosnes *et al.*<sup>(16)</sup> did observe improved (Sn 76%, Sp 80%,  $\kappa$  0.51) agreement when the NRS2002 screening component was removed from the GLIM.

To further the work being undertaken to validate the GLIM criteria in other patient groups and to examine whether GLIM is appropriate for use in vascular surgery units, the aim of this study was to determine the criterion validity of GLIM in diagnosing PEM in patients admitted to a vascular surgery unit using the PG-SGA as the comparator (Semi-gold standard<sup>(18)</sup>).

## Methods

This study is a retrospective analysis of baseline data collected during an observational study conducted from October 2014 to August 2016 that examined the nutritional status of adult patients admitted to a tertiary vascular surgery unit in Adelaide, South Australia. Data variables utilised in this study were chosen based on the recommendations outlined for validation of GLIM criteria<sup>(19)</sup>. A full description of the study and participant recruitment methods has been described elsewhere<sup>(1)</sup>. All patients over 18 years of age were eligible to participate but were excluded if they were admitted for day procedures only, were unable to be recruited within 72 h or were receiving palliative care. The study received ethical approval from the Southern Adelaide Health Research and Ethics Committee (approval number 258.14) and governance approval from the Flinders Medical Centre.

Within 72 h of admission, on entry to the study, the PG-SGA was conducted by an accredited practicing dietitian according to the methods of Ottery *et al.*<sup>(15)</sup>. Each participant was awarded a PG-SGA rating of PG-SGA A (well nourished), B (moderately or suspected malnutrition) or C (severely malnourished).

Retrospective determination of PEM according to the GLIM was completed using baseline parameters. Participants were only included in the analyses if they had all relevant parameters collected at baseline. The GLIM framework incorporates a validated screening tool of choice as the first step; however, in the current study, this step was not included as there was not a valid

screening tool completed at the time of data collection. For the phenotypic criteria, percentage weight loss was determined using self-reported weight history at 6 months prior to data collection, or where 6-month data were not able to be reported 1 month data were utilised. This was then compared with current weight to derive percentage loss over 6 months or 1 month, respectively. Body weight was collected using a calibrated weigh chair (HVL-CS Hospital Chair Scale, A&D Mercury Pty Ltd) to the nearest 0.1 kg. BMI was estimated using actual body weight and estimated height from ulna length<sup>(20)</sup>. Low BMI for age was determined as per the GLIM framework<sup>(9)</sup>. Muscle mass was determined using the Lunar Prodigy Pro dual-energy X-ray absorptiometer (DEXA) in conjunction with Encore software version 7.5. Appendicular skeletal muscle was calculated as the sum of the appendicular lean soft tissue in both upper and lower limbs and converted to appendicular skeletal muscle index by dividing the appendicular skeletal muscle mass by height squared (ASMI, kg/m<sup>2</sup>). Participants were classified as having low muscle mass if ASMI was <7.26 kg/m<sup>2</sup> in males and <5.25 kg/m<sup>2</sup> in females as per the GLIM framework<sup>(9)</sup>.

For the aetiologic criteria, information regarding reduction in food intake for 2 weeks or more was collected from the baseline PG-SGA along with data regarding the presence of gastrointestinal symptoms impacting food intake. Similarly, information regarding acute disease/injury or chronic disease-related inflammation was collected from baseline PG-SGA and medical case note entries. These variables included the presence of active liver, respiratory or renal disease, active cancer and/or blood malignancies, major abdominal surgery from the PG-SGA<sup>(15)</sup> as well as poorly controlled diabetes and medical diagnosis of inflammation in the case notes. Plasma C-reactive protein was measured according to the hospital laboratory and also utilised for the aetiologic criteria of inflammation if values were greater than 8.0 mg/l as per laboratory indicators.

Participants were diagnosed as malnourished according to the GLIM if they displayed at least one phenotypic and one aetiologic criterion as per the framework<sup>(9)</sup>.

## Statistical analysis

All analyses were conducted using SPSS for Windows version 27 (SPSS Inc.). Descriptive statistics were presented as mean and standard deviation or median (interquartile range) depending on normality. Sample characteristics were expressed as frequencies (*n*, %).

Diagnostic accuracy and consistency of the GLIM were examined. Sn, Sp, positive predictive value and negative predictive value were determined against the results of the PG-SGA (the reference standard) to determine the diagnostic accuracy of the GLIM in diagnosing patients with malnutrition according to recommendations<sup>(19)</sup>. As the PG-SGA results in three categories of nutritional status, PG-SGA B and PG-SGA C categories were amalgamated resulting in two categories of 'well nourished' and 'malnourished' to enable Sn and Sp analysis which is common practice in the literature<sup>(21–23)</sup>. The recommended cut points for Sn and Sp for determining diagnostic accuracy were set at 80% as per de van der Shueren *et al.*<sup>(19)</sup>.





Diagnostic consistency between the GLIM and PG-SGA was assessed using  $\kappa$  statistic. The value of  $\kappa$  varies from 0 to 1 with values <0.2 indicating poor, 0.21–0.4 fair, 0.41–0.6 moderate, 0.61–0.8 substantial and >0.8 as almost perfect concordance. Negative  $\kappa$  values indicate that the number of agreements observed is fewer than would be expected by chance indicating poor consistency overall<sup>(24)</sup>.

Due to the retrospective nature of this study and the inability to do a sample size calculation, a *post hoc* power calculation was performed to determine the statistical power of the findings. This was conducted using the prevalence of malnutrition found with the PG-SGA as the known population (comparator) and the prevalence of malnutrition found with the GLIM as the study group and an  $\alpha$  value of 0.05. *Post hoc* power calculation was conducted using the online ClinCalc *post hoc* power calculator to evaluate the statistical power of an existing study<sup>(25)</sup>.

**Results**

A total of 322 participants were recruited into the original study from a total of 902 eligible patients admitted to the vascular surgery unit<sup>(8)</sup>. Of the 322 participants, 224 had a full data set to enable determination of the GLIM and were included in this study.

Participant characteristics are shown in Table 1. The majority of participants were male (70.1 %) with a mean (SD) age of 67.3 (14.4) years and median (interquartile range) BMI of 27.8 (24.2, 32.3) kg/m<sup>2</sup>. Sixty-five (29 %) participants had at least one GLIM phenotype criterion and 194 (86.6 %) had at least one aetiological criterion. Overall, 64 (28.6 %) participants were classified as malnourished by the GLIM, and thirty-eight (17 %) by the PG-SGA.

Table 2 displays the diagnostic accuracy of the GLIM compared with the PG-SGA with Sn value of 73.7 % (95 % CI (52.8, 94.6)) and Sp 80.6 % (95 % CI (75.2, 86.0)) with negative predictive value of 93.8 % as well as positive predictive value of 43.8 %. There was an overlap of twenty-eight patients that were classified as malnourished by both methods (43.8 % of the GLIM malnourished and 73.7 % of the PG-SGA malnourished).  $\kappa$  was found to be 0.427 ( $P < 0.001$ ) indicating moderate diagnostic consistency.

**Table 1.** Participant characteristics (Numbers and percentage; mean values and standard deviations)

Characteristic	Participants	
	n	%
Sex	Males 157 Females 67	70.1 29.9
Age (years)		
Mean	67.3	
SD	14.4	
BMI (kg/m <sup>2</sup> )		
Median	27.8	
IQR	24.2, 32.3	
Malnourished (GLIM)	64	28.6
Malnourished (PG-SGA B and C)	38	17

IQR, interquartile range; GLIM, Global Leadership Initiative on Malnutrition; PG-SGA, Patient-Generated Subjective Global Assessment.

**Table 2.** Concurrent validity of the GLIM in a sample of 224 adult inpatients of a vascular surgery unit compared with the PG-SGA (Percentages and 95 % confidence intervals)

	%	95 % CI
Sensitivity	73.7 %	52.8, 94.6
Specificity	80.6 %	75.2, 86.0
Positive predictive value	43.8 %	
Negative predictive value	93.8 %	
$\kappa$ (P value)	0.427	<0.0001

GLIM, Global Leadership Initiative on Malnutrition; PG-SGA, Patient-Generated Subjective Global Assessment.

$$\text{Power} = \Phi \left\{ \frac{\sqrt{N * \frac{(P_1 - P_0)^2}{(P_0 * Q_0)} - Z_{1-\alpha/2}}}{\sqrt{\frac{P_1 * Q_1}{(P_0 * Q_0)}}} \right\}$$

$$\text{Power} = \Phi \left\{ \frac{\sqrt{224 * \frac{(0.286 - 0.18)^2}{(0.18 * 0.82)} - 1.96}}{\sqrt{\frac{0.286 * 0.714}{(0.18 * Q_0)}}} \right\}$$

Power =  $\Phi$  (1.844) = 0.967 = 96.7% Power

$P_0$  = proportion (incidence) of malnutrition according to PG-SGA

$P_1$  = proportion (incidence) of malnutrition according to GLIM

N = Sample size for the study

$\alpha$  = probability of type 1 error

z = critical Z value for a given  $\alpha$

$\Phi$ ( ) = function converting a critical Z value to power

**Fig. 1.** *Post hoc* power calculations. GLIM, Global Leadership Initiative on Malnutrition; PG-SGA, Patient-Generated Subjective Global Assessment.

The *post hoc* power calculation is shown in Fig. 1. Using an incidence of 17 % in the population (PG-SGA, comparator) and 28.6 % incidence in the study group (GLIM), and an  $\alpha$  of 0.05, a sample size of 224 resulted in a *post hoc* power of 98.7 %.

**Discussion**

This study adds further to research already conducted exploring the validity of GLIM in the diagnosis of PEM across clinical specialties. It also adds to the research examining the assessment of PEM in patients within vascular surgery units.

In patients within a vascular surgery unit, the GLIM reached a Sn of 73.7 % and Sp of 80.6 % which is approaching the cut-off value of 80 % that indicates a valid instrument. However, the positive predictive value was low indicating that whilst the GLIM was able to identify the same malnourished patients as the PG-SGA, it also has a high likelihood of over-diagnosing

malnutrition and hence may not be a valid assessment method when compared with the PG-SGA in this patient group.

Previous research examining the validity of the GLIM has produced Sn and Sp values of 43–85% and 69–79%, respectively<sup>(10–13,16,17)</sup>; however, the patient groups are varied, and the reference standards also differ across the studies making comparisons more challenging. Other differences can be observed across studies in how the presence of low muscle mass has been determined. Two studies<sup>(16,17)</sup> have utilised bio-electrical impedance assay to determine low muscle mass with another study<sup>(12)</sup> relying on hand-grip strength in addition to low BMI as an alternative method for muscle mass. Previous studies have utilised bio-electrical impedance assay to determine fat free mass (FFM), which can be affected by hydration status and less reliable in obese individuals and in PEM<sup>(26,27)</sup>. In the current study, FFM was determined using DEXA which is a preferred method and could be viewed as a more robust method compared with the other studies<sup>(28)</sup>.

Overall, the GLIM identified a higher proportion of patients as malnourished compared with the PG-SGA which may be due to the differences between the two methods. The PG-SGA incorporates subjective assessment of body composition as opposed to the objective methods used in the GLIM which could lead to under-estimation of muscle and fat depletion by the assessor. Objective measures of muscle stores in the GLIM eliminate the potential assessment bias associated with subjective measures. Another potential reason for the differences is the contribution of the different parameters to the overall diagnosis of nutritional status in the two methods. In the PG-SGA, the physical exam, nutrition impact symptoms and other parameters contribute different weightings to the overall assessment, whereas each criterion in the GLIM is of equal weighting to the overall assessment. Differences in the time frame of reported weight loss (1 month in the PG-SGA and 6 months in the GLIM) could also impact on differences in the overall diagnosis of malnutrition using both methods.

Overall prevalence of PEM in the participants of the current study was 17% (PG-SGA) and 28.6% (GLIM) which is lower than other studies examining patients in the vascular surgery setting<sup>(1,3,4)</sup>. However, it is dependent on the type of nutritional deficits being included in the assessment and the method of assessment employed. Whilst PEM is relevant in this patient group, micronutrient deficiencies are also relevant and prevalent<sup>(1)</sup> and are not captured with either assessment method examined in the current study.

In this study, only nineteen (8.5%) participants reported a weight loss of 5% or more and only 12.5% and 13.5% were found to have a low BMI or reduced muscle mass, respectively, so only 29% ( $n$  65) displayed the minimum of one phenotypic criterion required for the GLIM. Conversely, a high proportion (86.6%) of participants had at least one aetiological criterion, with 174 (77.65%) displaying the inflammation criterion and 102 (45.5%) reporting a reduced oral intake. These figures indicate that whilst patients in vascular surgery units may have reduced intake and/or inflammation, it is not translating to the traditional phenotypic criteria included in the GLIM and traditional measures of nutritional status that are incorporated in most assessment tools. Hence, to fully capture the extent of nutritional

deficits (PEM as well as micronutrients), an assessment tool incorporating both markers of PEM and micronutrients would be of great value.

When discussing the results, it is important to consider the strengths and limitations of the study. A key strength of the study is that muscle mass was determined using DEXA which is an objective, reliable method of determining muscle quantity according to the revised European consensus on the diagnosis of sarcopenia<sup>(28)</sup>. The *post hoc* power calculation demonstrated that the sample size was adequately powered to detect a difference between the two methods; hence, the results are not due to type 1 error. In addition, the sample size is still comparable or larger than those found in other tool validation studies<sup>(14,29–31)</sup> and well-cited recommendations<sup>(32)</sup>. A potential limitation to acknowledge is in the determination of data to address the phenotypic criterion pertaining to weight loss in the GLIM. In the GLIM, the cut-off used is >5% within the past 6 months (or >10% beyond 6 months). As these data were obtained from the PG-SGA collected in the original study, we were not able to determine whether the >5% weight loss reported in the PG-SGA was within 6 months or 1 month in some participants. Whilst there is an element of lack of clarity regarding this data, it would not affect the results as weight loss over 1 month or 6 months would meet the GLIM criterion. Another potential limitation could be the omission of the initial screening component of the GLIM which has also occurred in other validation studies<sup>(10,33)</sup>. Whilst the authors speculate that it is unlikely to it have affected the proportion diagnosed as malnourished in the subsequent assessment phase of the GLIM, future studies should examine the full GLIM including the screening component.

In conclusion, the GLIM framework for diagnosing malnutrition did not perform adequately in a cohort of patients admitted to a vascular surgery. A key nutritional issue in patients within vascular surgery settings is micronutrient-deficit, and hence the addition of parameters to identify these deficits in addition to PEM would be of great value.

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There are no conflicts of interest.

### References

1. Thomas J, Delaney C, Suen J, *et al.* (2019) Nutritional status of patients admitted to a metropolitan tertiary care vascular surgery unit. *Asia Pac J Clin Nutr* **28**, 64–70.



2. Durkin MT, Mercer KG, McNulty MF, *et al.* (1999) Vascular surgical society of Great Britain and Ireland: contribution of malnutrition to postoperative morbidity in vascular surgical patients. *Br J Surg* **86**, 702.
3. De Waele E, Moerman L, Van Bael K, *et al.* (2014) High incidence of malnutrition in elective vascular surgery patients: an observational auditing study. *J Translat Intern Med* **2**, 32–35.
4. Zhang SS, Tang ZY, Fang P, *et al.* (2013) Nutritional status deteriorates as the severity of diabetic foot ulcers increases and independently associates with prognosis. *Exp Ther Med* **5**, 215–222.
5. Westvik TS, Krause LK, Pradhan S, *et al.* (2006) Malnutrition after vascular surgery: are patients with chronic renal failure at increased risk? *Am J Surg* **192**, e22–e27.
6. Ambler G, Brooks D, Al Zuhir N, *et al.* (2015) Effect of frailty on short-and mid-term outcomes in vascular surgery patients. *Br J Surg* **102**, 638–645.
7. Gau BR, Chen HY, Hung SY, *et al.* (2016) The impact of nutritional status on treatment outcomes of patients with limb-threatening diabetic foot ulcers. *J Diabetes Complications* **30**, 138–142.
8. Thomas J, Kaambwa B, Delaney C, *et al.* (2019) An evaluation of the validity of nutrition screening and assessment tools in patients admitted to a vascular surgery unit. *Br J Nutr* **122**, 689–697.
9. Cederholm T, Jensen G, Correia M, *et al.* (2019) GLIM criteria for the diagnosis of malnutrition – a consensus report from the global clinical nutrition community. *J Cachexia Sarcopenia Muscle* **10**, 207–217.
10. Allard J, Keller H, Gramlich L, *et al.* (2020) GLIM criteria has fair sensitivity and specificity for diagnosing malnutrition when using SGA as comparator. *Clin Nutr* **39**, 2771–2777.
11. Clark AB, Reijnierse EM, Lim WK, *et al.* (2020) Prevalence of malnutrition comparing the GLIM criteria, ESPEN definition and MST malnutrition risk in geriatric rehabilitation patients: RESORT. *Clin Nutr* **39**, 3504–3511.
12. De Groot L, Lee G, Ackerie A, *et al.* (2020) Malnutrition screening and assessment in the cancer care ambulatory setting: mortality predictability and validity of the patient-generated subjective global assessment short form (PG-SGA SF) and the GLIM criteria. *Nutrients* **12**, 2287.
13. Theilla M, Rattanachaiwong S, Kagan I, *et al.* (2021) Validation of GLIM malnutrition criteria for diagnosis of malnutrition in ICU patients: an observational study. *Clin Nutr* **40**, 3578–3584.
14. Henrique JR, Pereira RG, Ferreira RS, *et al.* (2020) Pilot study GLIM criteria for categorization of a malnutrition diagnosis of patients undergoing elective gastrointestinal operations: a pilot study of applicability and validation. *Nutrition* **79–80**, 110961.
15. Ottery F (2000) Patient-generated subjective global assessment. In *The Clinical Guide to Oncology Nutrition*, pp. 11–23 [P McCallum and C Polisena, editors]. Chicago: The American Dietetic Association.
16. Rosnes KS, Henriksen C, Høidalen A, *et al.* (2021) Agreement between the GLIM criteria and PG-SGA in a mixed patient population at a nutrition outpatient clinic. *Clin Nutr* **40**, 5030–5037.
17. Ijmker-Hemink V, Heerschop S, Wanten G, *et al.* (2021) Evaluation of the validity and feasibility of the GLIM criteria compared with PG-SGA to diagnose malnutrition in relation to 1-year mortality in hospitalized patients. *J Acad Nutr Diet* **122**, 595–601.
18. Cederholm T & Barazzoni R (2021) A year with the GLIM diagnosis of malnutrition – does it work for older persons? *Curr Opin Clin Nutr Metab Care* **24**, 4–9.
19. de van der Schueren MAE, Keller H, Cederholm T, *et al.* (2020) Global leadership initiative on malnutrition (GLIM): guidance on validation of the operational criteria for the diagnosis of protein-energy malnutrition in adults. *Clin Nutr* **39**, 2872–2880.
20. BAPEN (2003) *The 'MUST' Explanatory Booklet. A Guide to the 'Malnutrition Universal Screening Tool' ('MUST') for Adults. Malnutrition Advisory Group (MAG): A Standing Committee of the British Association for Parenteral and Enteral Nutrition (BAPEN)*. Redditch, UK: BAPEN.
21. Chi J, Yin S, Zhu Y, *et al.* (2017) A comparison of the nutritional risk screening 2002 tool with the subjective global assessment tool to detect nutritional status in Chinese patients undergoing surgery with gastrointestinal cancer. *Gastroenterol Nurs* **40**, 19–25.
22. Stratton R, Hackston A, Longmore D, *et al.* (2004) Malnutrition in hospital outpatients and inpatients: prevalence, concurrent validity and ease of use of the 'malnutrition universal screening tool' ('MUST') for adults. *Br J Nutr* **92**, 799–808.
23. Donini LM, Poggiogalle E, Molfino A, *et al.* (2016) Mini-nutritional assessment, malnutrition universal screening tool, and nutrition risk screening tool for the nutritional evaluation of older nursing home residents. *J Am Med Dir Assoc* **17**, 959.e11–959.e18.
24. Landis J & Koch G (1977) The measurement of observer agreement for categorical data. *Biometrics* **33**, 159–174.
25. Kane S (2018) Post-Hoc Power Calculator. Evaluate Statistical Power of an Existing Study. <https://clincalc.com/stats/Power.aspx> (accessed February 2022).
26. Coppini LZ, Waitzberg DL & Campos ACL (2005) Limitations and validation of bioelectrical impedance analysis in morbidly obese patients. *Curr Opin Clin Nutr Metab Care* **8**, 329–332.
27. Mialich M, Faccioli Sicchieri J & Jordao Junior A (2014) Analysis of body composition: a critical review of the use of bioelectrical impedance analysis. *Int J Clin Nutr* **2**, 1–10.
28. Cruz-Jentoft AJ, Bahat G, Bauer J, *et al.* (2019) Sarcopenia: revised European consensus on definition and diagnosis. *Age Ageing* **48**, 16–31.
29. Pars H, Açıköz A & Erdoğan BD (2020) Validity and reliability of the Turkish version of three screening tools (PYMS, STAMP, and STRONG-kids) in hospitalized children. *Clin Nutr ESPEN* **39**, 96–103.
30. Wester P, Angus R, Easlea D, *et al.* (2018) Use of the malnutrition screening tool by non-dietitians to identify at-risk patients in a rehabilitation setting: a validation study. *Nutr Diet* **75**, 324–330.
31. Jackson HS, MacLaughlin HL, Vidal-Diez A, *et al.* (2019) A new renal inpatient nutrition screening tool (Renal iNUT): a multicenter validation study. *Clin Nutr* **38**, 2297–2303.
32. Willet W (1998) *Nutritional Epidemiology*, 2nd ed. New York: Oxford University Press.
33. López-Gómez JJ, Ballesteros-Pomar MD, Torres-Torres B, *et al.* (2021) Malnutrition at diagnosis in amyotrophic lateral sclerosis (ALS) and its influence on survival: using glim criteria. *Clin Nutr* **40**, 237–244.