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Cardiovascular and metabolic risk of antipsychotics in children and young adults: a multinational self-controlled case series study

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Abstract

Aims. The risk of antipsychotic-associated cardiovascular and metabolic events may differ among countries, and limited real-world evidence has been available comparing the corresponding risks among children and young adults. We, therefore, evaluated the risks of cardiovascular and metabolic events in children and young adults receiving antipsychotics.

Methods. We conducted a multinational self-controlled case series (SCCS) study and included patients aged 6–30 years old who had both exposure to antipsychotics and study outcomes from four nationwide databases of Taiwan (2004–2012), Korea (2010–2016), Hong Kong (2001–2014) and the UK (1997–2016) that covers a total of approximately 100 million individuals. We investigated three antipsychotics exposure windows (i.e., 90 days pre-exposure, 1–30 days, 30–90 days and 90 + days of exposure). The outcomes were cardiovascular events (stroke, ischaemic heart disease and acute myocardial infarction), or metabolic events (hypertension, type 2 diabetes mellitus and dyslipidaemia).

Results. We included a total of 48 515 individuals in the SCCS analysis. We found an increased risk of metabolic events only in the risk window with more than 90-day exposure, with a pooled IRR of 1.29 (95% CI 1.20–1.38). The pooled IRR was 0.98 (0.90–1.06) for 1–30 days and 0.88 (0.76–1.02) for 31–90 days. We found no association in any exposure window for cardiovascular events. The pooled IRR was 1.86 (0.74–4.64) for 1–30 days, 1.35 (0.74–2.47) for 31–90 days and 1.29 (0.98–1.70) for 90 + days.

Conclusions. Long-term exposure to antipsychotics was associated with an increased risk of metabolic events but did not trigger cardiovascular events in children and young adults.

Introduction

The average prevalence of psychiatric disorders was about 22.1% with the severe disorders (schizophrenia, bipolar disorder, severe depression, severe anxiety and severe post-traumatic stress disorder) estimated to be 5.1% (Charlson *et al.*, 2019). In children, the prevalence has been reported to be about 6.7% but varies between different countries (Erskine *et al.*, 2017). The use of antipsychotics has increased over the years and has become one of the mainstays for the treatment of psychiatric disorders in children, despite lingering concerns over side effects (Harrison *et al.*, 2012; Lao *et al.*, 2017; Lee *et al.*, 2018). Specifically, in such young populations, antipsychotic medications may induce cardiovascular and metabolic abnormalities (such as obesity, hyperglycaemia, dyslipidaemia and diabetes mellitus) that could affect their physical, mental and social development (Hsu *et al.*, 2013). Moreover, some life-threatening cardiovascular side effects of antipsychotics such as stroke, ischaemic heart disease (IHD) and acute myocardial infarction (AMI) have also been reported (De Hert *et al.*, 2011). Meanwhile, increasing prescribing of antipsychotics is observed among children and young

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adults (Olfson *et al.*, 2015), which leads to great concern regarding the safety of antipsychotics (Pillay *et al.*, 2018). Currently available studies mainly focus on the outcomes of weight change or shifts in metabolic parameters over a short period of time (Sjo *et al.*, 2017; Vandenberghe *et al.*, 2018). Not much is known about the specific risk of cardiovascular and metabolic events in children and young adults treated with antipsychotics, as previous studies were mainly of small sample size and thus may not have developed sufficient statistical power in their analyses (McIntyre and Jerrell, 2008; Burcu *et al.*, 2018).

The risk of antipsychotic-associated cardiovascular and metabolic events may differ among countries (Man et al., 2020). This could potentially be due to variations in healthcare systems and preferences of patients and clinicians in the choice of antipsychotics (Pillinger et al., 2020). From the genetic perspective, the variant HTR2C gene, which encodes the 5HT2c receptor, may increase metabolic risk, especially by the rs 1414334 C allele (Ma et al., 2014). A survey indicated that this allele is present in a higher proportion of Americans (10%) and Europeans (15%), but, by contrast, is of very low frequency in Asians (1%). This suggests that variations between different ethnicities could affect the risk of cardiovascular and metabolic events (Mulder et al., 2009). To date, only limited evidence has been available comparing the corresponding risks among children and young adults receiving antipsychotics between different populations in real-world situations. We, therefore, conducted the current study with four large population-based datasets from Taiwan, Korea, Hong Kong and the UK, which have coverage of approximately 100 million individuals in total, to evaluate and benchmark the risk of specific cardiovascular events (stroke, IHD and AMI) and metabolic events (hypertension, T2DM and dyslipidaemia), associated with antipsychotics in children and young adults.

Method

Database sources

We included databases from Taiwan (Taiwan's National Health Insurance Database; NHID), Korea (Korea's NHID), Hong Kong (Clinical Data Analysis and Reporting System; CDARS) and the UK (The Health Improvement Network; THIN) in this study (Hsieh et al., 2019; Ilomäki et al., 2020). Additional details about the included databases are presented in Supplementary Table 1. We applied a distributed network approach with a common data model (CDM) (Lai et al., 2015). Briefly, the coordination centre distributed the common SAS program for analysis, generating aggregated results based on the CDM that standardised the structures and contents of participating databases, and collected the final summary of results from each participating site (Lai et al., 2015). This approach preserved the confidentiality of the data because the raw data stayed in the local site while the analyses were executed respectively by the sites (Lai et al., 2018a). Moreover, we could maintain the consistency of analysis among the sites through the use of the common analysis program (Lai et al., 2018b). The details of the mapping codes for diagnosis and the CDM are presented in Supplementary Table 2 and Supplementary Table 3. The study has been approved by the Human Research Ethics Committee at National Cheng Kung University (No. NCKU HREC-E-105-259-2); the University of Hong Kong/Hospital Authority Hong Kong West Cluster (No. UW15-619); The Health Improvement Network Scientific Review Committee (19THIN087) and the Institutional Review Board of Sungkyunkwan University (SKKU 2018-04-106).

Study design

We applied a self-controlled case series (SCCS) design to investigate the association between antipsychotics and risk of metabolic/ cardiovascular event, whereby the relative risk is estimated based on within-person comparisons rather than between-person comparisons (Hallas and Pottegård, 2014; Petersen *et al.*, 2016). As a result, both measured and unmeasured time-independent confounding factors, such as sex, ethnicity, environmental and cultural factors are eliminated. The design is especially important for multinational studies, benchmarking results from different countries with very heterogeneous healthcare conditions (Lai *et al.*, 2018*a*).

Source population and exposure

We included patients aged 6–30 years diagnosed with a mental disorder (ICD-9-CM: 290–319) who were newly receiving oral antipsychotic drugs between 2004 and 2012 in Taiwan, 2010 and 2016 in Korea, 2001 and 2014 in Hong Kong and 1997 and 2016 in the UK. Incident use of antipsychotics was captured based on a 1-year washout period before the first record of antipsychotic prescription in the database. We excluded patients who had a record of congenital disorders, including congenital heart disease, familial hypercholesterolaemia and type 1 diabetes. We also excluded patients who had a cancer diagnosis record. Details of diagnostic codes are presented in Supplementary Table 3.

Case identification and ascertainments

We included patients who had a record of the outcome of interest (cardiovascular events included AMI, IHD and stroke, and metabolic events included hypertension, dyslipidaemia and T2DM) for the analyses. To improve the validity of diagnoses of metabolic events, we confirmed cases by the records for corresponding drug prescriptions, including oral hypoglycaemic agents (A10B) for T2DM, lipid-modifying agents (C10) for dyslipidaemia and antihypertensive drugs including diuretics (C03A and C03B), beta-blockers (C07, except propranolol; C07AA05), calcium channel blockers (C08CA) and angiotensin-converting enzyme inhibitors (ACEI) / angiotensin receptor blockers (ARB) (C09).

Definition of risk periods

We defined the observation period based on the availability of data sources as mentioned in the previous section. Observational periods began on the first available date in the corresponding database, or the sixth birthday of the patient (whichever was later) and ended on the last available date in the corresponding database, the 31st birthday of the patient, or registered date of death (whichever was earlier). For each included participant, we defined the risk periods with respect to the duration of exposure to antipsychotics and categorised them into five mutually exclusive windows: (1) 90 days before (pre-exposure), (2) 1-30 days, (3) 31-90 days, (4) more than 90 days of antipsychotics use and (5) 30-day after the end of antipsychotics use (post-exposure) (Fig. 1). Details of ATC codes for antipsychotics are presented in Supplementary Table 4. A pre-exposure period was added to take account of the possibility that the outcome of interest may affect the likelihood of antipsychotics treatment, which in turn may introduce bias into the risk estimate during



Fig. 1. Schematic presentation of self-control case series.

Table 1. Baseline characteristics of all antipsychotic users included in the analysis

| | Taiwan | Korea | Hong Kong | The UK |
|---------------------|---------------|----------------|-------------|-------------|
| Number of patients | 107 425 | 284 843 | 19 034 | 7770 |
| Sex, n (%) | | | | |
| Male | 65 532 (61.0) | 169 856 (59.6) | 9711 (51.0) | 4956 (63.8) |
| Female | 41 893 (39.0) | 114 987 (40.4) | 9323 (49.0) | 2814 (36.2) |
| Age (mean ± SD) | 21.1 (±6.7) | 19.9 (±6.9) | 23.3 (±5.3) | 24.9 (±3.8) |
| Age subgroup, n (%) | | | | |
| 6–12 years | 16 091 (15.0) | 50 829 (17.8) | 667 (3.50) | 7 (0.1) |
| 13–18 years | 18 986 (17.7) | 67 759 (23.8) | 3048 (16.0) | 355 (4.6) |
| 19–24 years | 32 837 (30.6) | 80 726 (28.3) | 6451 (33.9) | 3180 (40.9) |
| 25–30 years | 39 511 (36.8) | 85 529 (30.3) | 8868 (46.6) | 4228 (54.4) |

treatment whereas the post-exposure period acts as a washout period. To manage possible confounding effects due to age, we performed 1-year age banding for all patients in the analysis (Petersen *et al.*, 2016).

Statistical analysis

We report characteristics of patients included in the analyses by countries and describe categorical variables (e.g., sex) by number with proportion and continuous variables (e.g., age) by mean with standard deviation (SD). We considered the entire study period and categorised the risk periods in a time-varying manner, i.e. the exposure status was updated according to the treatment time. We calculate incidence rate ratio (IRR) and 95% confidence interval (95% CI) by conditional Poisson regression, adjusted for age in 1-year age bands to evaluate the risk of specific cardiovascular and metabolic events associated with antipsychotics in the different risk windows (Petersen *et al.*, 2016). IRRs for each risk window in each site were pooled by the random-effect model (DerSimonian and Laird, 1986). The Taiwanese and Korean databases provided a sufficient number of cases to conduct a secondary analysis for the risk comparisons between individual antipsychotics.

Results

Antipsychotic users from each country

We identified a total of 107 425 patients from Taiwan, 284 843 from Korea, 19 034 patients from Hong Kong and 7770 patients from the UK. The mean age $(\pm SD)$ upon receiving the first

prescription was 21.1 ± 6.7 years in Taiwan, 23.3 ± 5.3 years in Hong Kong and 24.9 ± 3.8 years in the UK. We found more males than females in Taiwan (61%), Korea (60%) and the UK (64%) and about the same males and females in Hong Kong (Table 1). The patterns and rates of antipsychotics use are presented in Fig. 2. The proportion of antipsychotics use varied among countries. The most commonly prescribed drugs at initiation were sulpiride in Taiwan, risperidone in Korea, haloperidol in Hong Kong and olanzapine in the UK.

Pooled analysis

There were a total of 9730 (overall incidence rate, 11.5 per 100 person-years), 38 432 (12.4 per 100 person-years), 233 (7.6 per 100 person-years) and 120 (11.9 per 100 person-years) patients with at least one record of cardiovascular or metabolic events in Taiwan, Korea, Hong Kong and the UK, respectively. An increased risk of metabolic events was observed with more than 90-days of antipsychotics exposure with the pooled IRR of 1.29 (95% CI 1.20-1.38), but not in other risk windows. The corresponding IRR was 0.98 (95% CI 0.90-1.06) for 1-30 days and 0.88 (0.76-1.02) for 31-90 days of antipsychotics exposure. The I^2 for heterogeneity ranged from 0% to 39.5% in the pooled estimates from metabolic events. For cardiovascular events, no significant association was identified in any risk windows in the pooled analysis. The pooled IRR was 1.86 (0.74-4.64) for 1-30 days, 1.35 (0.74-2.47) for 31-90 days and 1.29 (0.98-1.70) for more than 90-days of antipsychotics exposure, however, with high heterogeneity (I^2 ranged from 82.4% to 98.7%). No



Fig. 2. Utilisation pattern of antipsychotics among countries: (a) Taiwan (b) Korea (c) Hong Kong (d) UK.

significant associations were found in the pre- and post-exposure windows for both outcomes (Figs 3 and 4).

Analysis by sites and specific events

The results of SCCS from individual sites varied. We found the risks of metabolic events were higher in the risk period of 90 + days in Taiwan and Korea. The IRR were 1.27 (1.12–1.45) and 1.48 (1.38–1.59) for hypertension, 1.35 (1.12–1.64) and 1.38 (1.28–1.49) for T2DM and 1.36 (1.20–1.55) and 1.36 (1.28–1.49) for dyslipidaemia in Taiwan and Korea, respectively. However, the risk of metabolic events in the period of 90 + days was higher but not reached statistical significance in Hong Kong and the UK (Table 2).



Fig. 2. Continued.

Compared to the non-exposure period, we found the risk of stroke to be higher in the pre-exposure period (5.49; 4.72-6.38), 1–30 days (3.51; 3.07-4.01), 31–90 days (2.06; 1.75-2.42) and 90 + days (2.07; 1.77-2.43) and the risk of IHD (1.59; 1.35-1.86) and AMI was higher in the pre-exposure period (6.42;

2.86–14.45) and the risk period of 1–30 days (2.29; 0.91–5.78) in Taiwan (Table 3). We did not find an association between antipsychotics and cardiovascular events in Korea, except for a higher risk of stroke in the period of 90 + days (1.38; 1.34–1.41) comparing with the non-exposure period. In Hong Kong, we found the

| Risk window and Site | IRR (95% CI) | % Weigh |
|---|-------------------|------------|
| | | |
| Taiwan | 2 52 (2 25 2 82) | 7 10 |
| Hong Kong | 1 40 (0 56 3 52) | 3.01 |
| Korea | 0.08 (0.04, 0.19) | 3.61 |
| | 2.04 (0.74, 5.66) | 2.65 |
| Subgroup, DL (1 ² = 96.0%, p = 0.000) | 0.88 (0.17, 4.39) | 16.46 |
| Risk (1-30 days) | | |
| Taiwan 🔶 | 1.02 (0.92, 1.14) | 7.21 |
| Hong_Kong | 1.16 (0.62, 2.17) | 4.40 |
| Korea 🔶 | 0.90 (0.78, 1.04) | 7.10 |
| UK | 0.79 (0.36, 1.74) | 3.56 |
| Subgroup, DL (1 ² = 0.0%, p = 0.486) | 0.98 (0.90, 1.06) | 22.27 |
| Risk (31-90 days) | | |
| Taiwan | 0.81 (0.72, 0.91) | 7.18 |
| Hong_Kong | 0.63 (0.28, 1.40) | 3.49 |
| Korea | 0.98 (0.85, 1.13) | 7.11 |
| UK | 1.04 (0.50, 2.16) | 3.84 |
| Subgroup, DL (I [*] = 39.5%, p = 0.175) | 0.88 (0.76, 1.02) | 21.6 |
| Risk (>90 days) | 4.04/4.00 4.47) | 7.00 |
| | 1.34 (1.23, 1.47) | 1.23 |
| Hong_kong | 0.99 (0.60, 1.66) | 5.05 |
| Korea | 1.21 (1.07, 1.37) | 1.10 |
| Subgroup $DL/l^2 = 0.0\%$ p = 0.429) | 1.27 (0.69, 2.33) | 24.0 |
| Subgroup, DL (1 = 0.0%, p = 0.423) | 1.29 (1.20, 1.30) | 24.0 |
| Post-risk (1-30 days) | | |
| Taiwan 🚽 | 0.87 (0.70, 1.07) | 6.83 |
| Hong_Kong | 0.53 (0.12, 2.27) | 1.50 |
| Korea | 0.95 (0.69, 1.29) | 6.30 |
| UK 2 | 0.34 (0.05, 2.57) | 0.9 |
| Subgroup, DL (l ² = 0.0%, p = 0.661) | 0.88 (0.74, 1.05) | 15.65 |
| Heterogeneity between groups: $p = 0.000$ | | |
| Overall, DL (I ⁻ = 93.9%, p = 0.000) | 0.98 (0.79, 1.20) | 100.00 |
| 1 1 | 100 | |
| NOTE: Weights and between-subgroup heterogeneity test are from random-effects m | odel | |

Fig. 3. Pooled estimates of risk in metabolic events.

risk of cardiovascular events, specifically for stroke (5.80; 1.91– 17.61) was higher in the risk period of 1–30 days compared to the non-exposure period. We did not find any association between the use of antipsychotics and the risk of cardiovascular events in the UK.

Discussion

Previous studies evaluating the safety of antipsychotics in young people were generally with a limited sample size to acquire a precise estimation (McIntyre and Jerrell, 2008; Burcu *et al.*, 2018). The current study used four large databases from Taiwan, Korea, Hong Kong and the UK to evaluate the risk of cardiovascular and metabolic events associated with antipsychotics in children, adolescents and young adults. We identified an increased risk of metabolic events with exposure to antipsychotics for more than 90 days from the pooled results. However, no significant association was found between the use of antipsychotics and the risk of cardiovascular events. This suggested that the increased risk in the metabolic events may not be severe enough to trigger cardiovascular events in children, adolescents and young adults. Despite the varied risk pattern among countries, the conclusion is consistent with the overall results.

The results of our study were largely consistent with previous studies regarding the association between the risk of metabolic events with antipsychotics use (Man *et al.*, 2020). However, we found an increased risk of metabolic events only with the antipsychotics exposure for more than 90 days, implicating it may require a period of accumulative exposure of drugs to develop metabolic events. Our pooled analysis did not support the use of antipsychotics was associated with cardiovascular events. However, the finding was based on the results from countries with high heterogeneity. We found an increased risk of stroke in Hong Kong in the first 30 days of antipsychotics treatment.

| Risk_window and Site | IRR (95% CI) | % Weight |
|---|----------------------|-------------|
| Pre-risk (-90 to 0 days) | | |
| Taiwan | 4.51 (4.03, 5.05) | 7.85 |
| Hong_Kong | 4.42 (1.17, 16.67) | 3.77 |
| Korea | 0.13 (0.08, 0.20) | 7.00 |
| Subgroup, DL (l ² = 99.1%, p = 0.000) | 1.34 (0.09, 20.19) | 18.62 |
| Risk (1-30 days) | | |
| Taiwan | 2.45 (2.21, 2.71) | 7.87 |
| Hong_Kong | 5.44 (2.10, 14.10) | 5.05 |
| Korea | 0.80 (0.72, 0.89) | 7.86 |
| UK | - 0.43 (0.01, 15.82) | 0.84 |
| Subgroup, DL (I ² = 98.7%, p = 0.000) | 1.86 (0.74, 4.64) | 21.62 |
| Risk (31-90 days) | | |
| Taiwan | 1.50 (1.32, 1.69) | 7.84 |
| Hong Kong | 3.73 (1.29, 10.80) | 4.64 |
| Korea | 0.77 (0.69, 0.86) | 7.86 |
| Subgroup, DL (I ² = 97.1%, p = 0.000) | 1.35 (0.74, 2.47) | 20.34 |
| Risk (>90 days) | | |
| Taiwan | 1.43 (1.27, 1.62) | 7.84 |
| Hong_Kong | 1.10 (0.30, 4.02) | 3.86 |
| Korea | 1.07 (0.98, 1.17) | 7.88 |
| UK H | 2.98 (1.00, 8.89) | 4.53 |
| Subgroup, DL (I ² = 82.4%, p = 0.001) | 1.29 (0.98, 1.70) | 24.11 |
| Post-risk (1-30 days) | | |
| Taiwan | 1.31 (1.02, 1.68) | 7.62 |
| Korea | 0.89 (0.72, 1.11) | 7.69 |
| Subgroup, DL (I ² = 81.0%, p = 0.022) | 1.07 (0.74, 1.57) | 15.31 |
| Heterogeneity between groups: p = 0.837 | | |
| Overall, DL (l ² = 98.4%, p = 0.000) | 1.40 (0.98, 1.99) | 100.00 |
| | 100 | |
| NOTE: Weights and between-subgroup heterogeneity test are from random-effects model | 100 | |

Fig. 4. Pooled estimates of risk in cardiovascular events.

In particular, we found patients receiving haloperidol had an increased risk of stroke in the initial stage of treatment in both Taiwan and Hong Kong. Haloperidol has a high affinity for binding to $\alpha 1$ and $\alpha 2$ receptors (Hensiek and Trimble, 2002), and the blockade of α receptors could cause fluctuations in blood pressure along with some symptoms such as hypotension, hypertension and QT interval prolongation, leading to a high risk of cardiovascular events (Cooper et al., 2016; Hiremath et al., 2019). However, a more parsimonious interpretation of this pattern of temporal association is that the observed increased risk of cardiovascular events is not due to antipsychotics but precedes it because an increased risk was also observed in 90 days pre-exposure period from both Taiwan and Hong Kong. The changes in behavioural and mental health symptoms or associated impairment that lead to a medical consultation or comorbidities, which in turn may contribute to the decision to prescribe antipsychotics.

Besides, it is noteworthy that different risk profiles among countries could be attributed to different patterns of antipsychotic or prescribing preferences among the countries. From the study cohorts, we found that SGA uses increased over the years in Korea and the UK, and until 2016, SGA accounted for more than 80% of total antipsychotics use in children, adolescents and young adults. The frequent use of SGA may explain the higher risk of metabolic events within the observation window of more than 90 days after drug initiation, compared to the non-exposure period. We suggest interpreting the results cautiously considering countries' specific situations. For instance, the healthcare accessibility or copayment of medical treatment may be different among countries, leading to various thresholds for seeking medical treatment (Lai et al., 2018b). The possibilities of the capture of events were different among countries. Moreover, the differences in the respective healthcare systems, cultures, behaviours of prescribing, preferences of clinicians among countries may also contribute to the heterogeneity of the results. Therefore we could not make inferences on the ethnic differences regarding the adverse effects in our study.

Table 2. Risk of metabolic events among countries

| | Taiwan NHID | | | Korea NHID | | | | | Hong Kong | | UK THIN | | | | | |
|-----------------------------|-------------|--------------|------|-------------|--------|--------------|------|-------------|-----------|--------------|---------|--------------|--------|--------------|------|--------------|
| | Events | Person-years | IRR | (95% CI) | Events | Person-years | IRR | (95% CI) | Events | Person-years | IRR | (95% CI) | Events | Person-years | IRR | (95% CI) |
| Metabolic risk | | | | | | | | | | | | | | | | |
| Baseline | 3898 | 37 852 | 1.00 | (reference) | 20 495 | 163 759 | 1.00 | (reference) | 89 | 1581 | 1.00 | (reference) | 44 | 511 | 1.00 | (reference) |
| Pre-risk (—90 to 0 days) | 352 | 1207 | 2.52 | (2.25–2.82) | 1763 | 23 396 | 0.08 | (0.04–0.19) | 5 | 40 | 1.40 | (0.56–3.52) | 6 | 19 | 2.04 | (0.74–5.66) |
| Risk (1–30 days) | 437 | 3935 | 1.02 | (0.92–1.14) | 1980 | 23 231 | 0.90 | (0.78–1.04) | 13 | 137 | 1.16 | (0.62–2.17) | 10 | 93 | 0.79 | (0.36–1.74) |
| Risk (31–90 days) | 375 | 3891 | 0.81 | (0.72–0.91) | 9333 | 54 134 | 0.98 | (0.85–1.13) | 7 | 136 | 0.63 | (0.28–1.40) | 13 | 92 | 1.04 | (0.50–2.16) |
| Risk (90 + days) | 1130 | 6825 | 1.34 | (1.23–1.47) | 61 | 5486 | 1.21 | (1.07–1.37) | 74 | 619 | 0.99 | (0.60-1.66) | 42 | 240 | 1.27 | (0.69–2.33) |
| Post-risk (1–30 days) | 90 | 838 | 0.87 | (0.70-1.07) | 447 | 3953 | 0.95 | (0.69–1.29) | 2 | 23 | 0.53 | (0.12–2.27) | 1 | 16 | 0.34 | (0.05–2.57) |
| Risk of hypertension | | | | | | | | | | | | | | | | |
| Baseline | 2038 | 19 906 | 1.00 | (reference) | 4886 | 39 465 | 1.00 | (reference) | 45 | 726 | 1.00 | (reference) | 9 | 92 | 1.00 | (reference) |
| Pre-risk (—90 to 0 days) | 201 | 623 | 2.81 | (2.42–3.26) | 11 | 1355 | 0.06 | (0.03-0.11) | 2 | 18 | 0.89 | (0.21–3.75) | 3 | 4 | 7.72 | (1.40-42.68) |
| Risk (1–30 days) | 251 | 1996 | 1.16 | (1.01–1.34) | 480 | 6172 | 0.61 | (0.56–0.68) | 9 | 55 | 1.41 | (0.64–3.11) | 1 | 16 | 0.57 | (0.06-5.14) |
| Risk (31–90 days) | 209 | 1974 | 0.90 | (0.77–1.05) | 587 | 6130 | 0.72 | (0.65–0.79) | 3 | 55 | 0.47 | (0.14–1.59) | 2 | 16 | 1.17 | (0.21-6.60) |
| Risk (90 + days) | 530 | 3252 | 1.27 | (1.12–1.45) | 2667 | 14 861 | 1.48 | (1.38–1.59) | 23 | 208 | 0.55 | (0.24–1.26) | 9 | 38 | 2.18 | (0.53–9.03) |
| Post-risk (1–30 days) | 38 | 427 | 0.72 | (0.52–1.00) | 118 | 1015 | 0.88 | (0.73–1.06) | 1 | 10 | 0.41 | (0.05-3.24) | 0 | 3 | N/A | () |
| Risk of type 2 DM | | | | | | | | | | | | | | | | |
| Baseline | 791 | 7618 | 1.00 | (reference) | 4154 | 33 428 | 1.00 | (reference) | 37 | 672 | 1.00 | (reference) | 19 | 208 | 1.00 | (reference) |
| Pre-risk (—90 to 0 days) | 64 | 254 | 2.13 | (1.64–2.77) | 11 | 1158 | 0.07 | (0.04–0.13) | 2 | 18 | 1.52 | (0.35–6.65) | 2 | 8 | 1.34 | (0.24–7.37) |
| Risk (1–30 days) | 85 | 867 | 0.88 | (0.69–1.11) | 422 | 5562 | 0.62 | (0.56–0.69) | 3 | 73 | 0.57 | (0.17–1.96) | 5 | 40 | 0.94 | (0.29–3.07) |
| Risk (31–90 days) | 78 | 857 | 0.79 | (0.61–1.01) | 389 | 5525 | 0.54 | (0.48-0.60) | 3 | 73 | 0.57 | (0.17–1.97) | 7 | 40 | 1.34 | (0.45–3.98) |
| Risk (90 + days) | 286 | 1629 | 1.35 | (1.12–1.64) | 2441 | 13 957 | 1.38 | (1.28–1.49) | 43 | 338 | 1.34 | (0.65–2.76) | 17 | 101 | 1.08 | (0.41–2.86) |
| Post-risk (1–30 days) | 22 | 184 | 0.86 | (0.56–1.33) | 106 | 887 | 0.84 | (0.69–1.03) | 0 | 12 | N/A | | 0 | 7 | N/A | () |
| Risk of dyslipidaemia | | | | | | | | | | | | | | | | |
| Baseline | 1773 | 17 368 | 1.00 | (reference) | 14 562 | 116 747 | 1.00 | (reference) | 12 | 316 | 1.00 | (reference) | 18 | 230 | 1.00 | (reference) |
| Pre-risk (—90 to 0 days) | 139 | 563 | 2.19 | (1.83–2.61) | 45 | 3885 | 0.09 | (0.06-0.12) | 1 | 7 | 2.33 | (0.27–20.11) | 1 | 8 | 0.92 | (0.08–10.32) |
| Risk (1–30 days) | 163 | 1908 | 0.78 | (0.66–0.92) | 1123 | 16 207 | 0.56 | (0.52–0.59) | 1 | 27 | 0.81 | (0.10-6.87) | 4 | 40 | 0.82 | (0.23–2.94) |
| Risk (31–90 days) | 165 | 1887 | 0.71 | (0.60-0.85) | 1292 | 16 091 | 0.64 | (0.61-0.68) | 1 | 27 | 0.82 | (0.10-6.90) | 4 | 39 | 0.82 | (0.23–2.95) |
| Risk (90 + days) | 613 | 3534 | 1.36 | (1.20–1.55) | 6376 | 37 150 | 1.38 | (1.34–1.42) | 21 | 160 | 2.59 | (0.77-8.67) | 19 | 110 | 1.43 | (0.56-3.63) |
| Post-risk (1–30 days) | 52 | 399 | 0.98 | (0.74–1.30) | 303 | 2771 | 0.88 | (0.78–1.00) | 1 | 4 | 2.69 | (0.25-29.41) | 1 | 6 | 0.81 | (0.10-6.73) |

NHID, National Health Insurance Database; CDARS, Clinical Data Analysis and Reporting System, THIN, The Health Improvement Network, DM, diabetes mellitus, N/A, not available.

Table 3. Risk of cardiovascular events among countries

| | | Taiwan NHID | | | | Korea NHID | | | | Hong K | ong CDARS | i | UK THIN | | | |
|-------------------------------|--------|------------------|------|--------------|--------|------------------|------|--------------|--------|------------------|-----------|---------------|---------|------------------|------|--------------|
| | Events | Person- years | IRR | (95% CI) | Events | Person- years | IRR | (95% CI) | Events | Person- years | IRR | (95% CI) | Events | Person- years | IRR | (95% CI) |
| Risk of cardiovascular events | | | | | | | | | | | | | | | | |
| Baseline | 2480 | 27 250 | 1.00 | (reference) | 4209 | 32 891 | 1.00 | (reference) | 17 | 402 | 1.00 | (reference) | 4 | 34 | 1.00 | (reference) |
| Pre-risk (–90 to 0 days) | 367 | 815 | 4.51 | (4.03–5.05) | 19 | 1101 | 0.13 | (0.08-0.20) | 3 | 9 | 4.42 | (1.17–16.67) | 0 | 1 | N/A | () |
| Risk (1–30 days) | 506 | 2384 | 2.45 | (2.21–2.71) | 443 | 4159 | 0.80 | (0.72-0.89) | 10 | 31 | 5.44 | (2.10-14.10) | 1 | 4 | 0.43 | (0.01-15.82) |
| Risk (31–90 days) | 324 | 2357 | 1.50 | (1.32–1.69) | 456 | 4128 | 0.77 | (0.69-0.86) | 7 | 31 | 3.73 | (1.29–10.80) | 0 | 4 | N/A | () |
| Risk (90 + days) | 467 | 3386 | 1.43 | (1.27–1.62) | 1231 | 8537 | 1.07 | (0.98-1.17) | 8 | 85 | 1.10 | (0.30-4.02) | 1 | 12 | 12 | (0.00-8.89) |
| Post-risk (1–30 days) | 65 | 531 | 1.31 | (1.02–1.68) | 86 | 749 | 0.89 | (0.72-1.11) | 0 | 6 | N/A | () | 0 | 1 | N/A | () |
| Risk of stroke | | | | | | | | | | | | | | | | |
| Baseline | 1175 | 14 308 | 1.00 | (reference) | 2042 | 16 312 | 1.00 | (reference) | 12 | 293 | 1.00 | (reference) | 4 | 31 | 1.00 | (reference) |
| Pre-risk (–90 to 0 days) | 212 | 434 | 5.49 | (4.72–6.38) | 6 | 547 | 0.09 | (0.07-0.11) | 2 | 7 | 4.29 | (0.87-21.23) | 0 | 1 | N/A | () |
| Risk (1–30 days) | 321 | 1253 | 3.51 | (3.07-4.01) | 226 | 2055 | 0.60 | (0.57–0.63) | 8 | 25 | 5.80 | (1.91–17.61) | 0 | 3 | N/A | () |
| Risk (31–90 days) | 196 | 1239 | 2.06 | (1.75-2.42) | 267 | 2040 | 0.68 | (0.65-0.71) | 6 | 24 | 4.25 | (1.26-14.26) | 0 | 3 | N/A | () |
| Risk (90 + days) | 284 | 1779 | 2.07 | (1.77–2.43) | 636 | 4310 | 1.38 | (1.34–1.41) | 6 | 69 | 1.01 | (0.21-4.76) | 1 | 10 | 0.51 | (0.01-20.82) |
| Post-risk (1–30 days) | 33 | 276 | 1.50 | (1.06-2.14) | 41 | 366 | 0.90 | (0.82–0.99) | 0 | 5 | N/A | () | 0 | 1 | N/A | () |
| Risk of IHD | | | | | | | | | | | | | | | | |
| Baseline | 1353 | 13 388 | 1.00 | (reference) | 2158 | 16 533 | 1.00 | (reference) | 3 | 75 | 1.00 | (reference) | 0 | 2 | 1.00 | (reference) |
| Pre-risk (–90 to 0 days) | 155 | 393 | 3.51 | (2.96–4.17) | 13 | 551 | 0.17 | (0.10-0.29) | 1 | 2 | 13.34 | (0.42-428.54) | 0 | 0.2 | N/A | () |
| Risk (1–30 days) | 196 | 1182 | 1.59 | (1.35–1.86) | 209 | 2079 | 0.71 | (0.61-0.82) | 1 | 5 | 9.42 | (0.34–258.63) | 1 | 0.3 | N/A | () |
| Risk (31–90 days) | 125 | 1169 | 0.95 | (0.77-1.15) | 196 | 2063 | 0.62 | (0.52- 0.72) | 1 | 5 | 9.42 | (0.34-258.63) | 0 | 0.3 | N/A | () |
| Risk (90 + days) | 196 | 1675 | 0.95 | (0.79–1.14) | 576 | 4151 | 0.94 | (0.83–1.07) | 1 | 14 | N/A | () | 0 | 0.3 | N/A | () |
| Post-risk (1–30 days) | 32 | 265 | 1.07 | (0.75–1.53) | 47 | 380 | 0.87 | (0.65–1.17) | 0 | 1 | N/A | () | 0 | 0.1 | N/A | () |
| Risk of AMI | | | | | | | | | | | | | | | | |
| Baseline | 35 | 372 | 1.00 | (reference) | 122 | 990 | 1.00 | (reference) | 2 | 34 | 1.00 | (reference) | 0 | 2 | 1.00 | (reference) |
| Pre-risk (–90 to 0 days) | 8 | 12 | 6.42 | (2.86–14.45) | 0 | 34 | N/A | () | 0 | 0.2 | N/A | () | 0 | 0.2 | N/A | () |
| Risk (1–30 days) | 6 | 34 | 2.29 | (0.91–5.78) | 19 | 127 | 1.22 | (0.71-2.10) | 1 | 2 | N/A | () | 0 | 0.4 | N/A | () |
| Risk (31–90 days) | 1 | 33 | 0.39 | (0.05–2.92) | 13 | 127 | 0.88 | (0.48–1.63) | 0 | 2 | N/A | () | 0 | 0.4 | N/A | () |
| Risk (90 + days) | 6 | 42 | 1.05 | (0.31–3.54) | 43 | 278 | 1.37 | (0.84–2.24) | 1 | 2 | N/A | () | 0 | 0.9 | N/A | () |
| Post-risk (1–30 days) | 0 | 8 | N/A | () | 1 | 23 | 0.46 | (0.06-3.31) | 0 | 0.4 | N/A | () | 0 | 0.1 | N/A | () |

NHID, National Health Insurance Database; CDARS, Clinical Data Analysis and Reporting System; THIN, The Health Improvement Network; IHD, ischaemic heart disease, AMI, acute myocardial infarction; N/A, not available.

Limitations

We were unable to assess the actual medication adherence of patients, which may cause a bias towards null because patients may or may not be taking the medication. The self-controlled design eliminated time-constant unmeasured confounders, but the results may be influenced by time-variant factors that were not associated with age (Hallas and Pottegård, 2014; Petersen *et al.*, 2016). Protopathic bias should be noted with patients who had undetected cardiovascular events that caused psychosis and the use of antipsychotics. Because clinicians may avoid antipsychotics which are well known to increase metabolic side effects, such as olanzapine for patients with higher baseline risk, the sample size and power of the study may be decreased. On the other hand, because metabolic monitoring is not being implemented routinely in all study countries, we may omit some patients with mild metabolic syndrome.

Conclusion

Exposure to antipsychotics for more than 90 days was associated with an increased risk of a metabolic event, but did not trigger cardiovascular events in children and young adults. Although we found varied risk profiles of cardiovascular and metabolic events between countries, the conclusion remained consistent with the overall results. Nevertheless, clinicians should be mindful of the possible cardiovascular and metabolic risk as with the use of all antipsychotics in children and young adults in clinical practice while long-term use of antipsychotics is required.

Supplementary material. The supplementary material for this article can be found at https://doi.org/10.1017/S2045796021000494.

Data. All data are not publicly available due to the issue of privacy. Requests for data require the review by the custodian from each country.

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Data analysis: Su CC and YC Chang. Interpretation of results: All authors. First draft of manuscript: ECC Lai.

Review and revise manuscript: All authors.

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Conflict of interest. None.

Ethical standards. The study has been approved by the Human Research Ethics Committee at National Cheng Kung University (No. NCKU HREC-E-105-259-2); the University of Hong Kong/Hospital Authority Hong Kong West Cluster (No. UW15-619); The Health Improvement Network Scientific Review Committee (19THIN087) and the Institutional Review Board of Sungkyunkwan University (SKKU 2018-04-106).

References

- Burcu M, Zito JM, Safer DJ, Magder LS, dosReis S, Shaya FT and Rosenthal GL (2018) Cardiovascular events following treatment initiation with atypical antipsychotic medications in publicly insured U.S. Youth. *Journal of Child and Adolescent Psychopharmacology* 28, 445–453.
- Charlson F, van Ommeren M, Flaxman A, Cornett J, Whiteford H and Saxena S (2019) New WHO prevalence estimates of mental disorders in

conflict settings: a systematic review and meta-analysis. Lancet 394, 240–248.

- **Cooper SJ, Reynolds GP, With expert co-authors (in alphabetical order)**, Barnes T, England E, Haddad PM, Heald A, Holt R, Lingford-Hughes A, Osborn D, McGowan O, Patel MX, Paton C, Reid P, Shiers D and Smith J (2016) BAP Guidelines on the management of weight gain, metabolic disturbances and cardiovascular risk associated with psychosis and antipsychotic drug treatment. Journal of Psychopharmacology 30, 717–748.
- De Hert M, Detraux J, van Winkel R, Yu W and Correll CU (2011) Metabolic and cardiovascular adverse effects associated with antipsychotic drugs. *Nature Reviews Endocrinology* **8**, 114–126.
- DerSimonian R and Laird N (1986) Meta-analysis in clinical trials. Controlled Clinical Trials 7, 177–188.
- Erskine HE, Baxter AJ, Patton G, Moffitt TE, Patel V, Whiteford HA and Scott JG (2017) The global coverage of prevalence data for mental disorders in children and adolescents. *Epidemiology and Psychiatric Sciences* 26, 395–402.
- Hallas J and Pottegård A (2014) Use of self-controlled designs in pharmacoepidemiology. *Journal of Internal Medicine* 275, 581–589.
- Harrison JN, Cluxton-Keller F and Gross D (2012) Antipsychotic medication prescribing trends in children and adolescents. *Journal of Pediatric Health Care* 26, 139–145.
- Hensiek AE and Trimble MR (2002) Relevance of new psychotropic drugs for the neurologist. *Journal of Neurology, Neurosurgery, and Psychiatry* 72, 281–285.
- Hiremath S, Ruzicka M, Petrcich W, McCallum MK, Hundemer GL, Tanuseputro P, Manuel D, Burns K, Edwards C, Bugeja A, Magner P, McCormick B, Garg AX, Rhodes E and Sood MM (2019) Alpha-blocker use and the risk of hypotension and hypotension-related clinical events in women of advanced age. *Hypertension* 74, 645–651.
- Hsieh CY, Su CC, Shao SC, Sung SF, Lin SJ, Kao Yang YH and Lai EC (2019) Taiwan's national health insurance research database: past and future. *Clinical Epidemiology* 11, 349–358.
- Hsu YC, Chien IC, Tan HK, Lin CH, Cheng SW, Chou YJ and Chou P (2013) Trends, correlates, and disease patterns of antipsychotic use among children and adolescents in Taiwan. *Social Psychiatry and Psychiatric Epidemiology* **48**, 1889–1896.
- Ilomäki J, Bell JS, Chan AYL, Tolppanen AM, Luo H, Wei L, Lai EC, Shin JY, De Paoli G, Pajouheshnia R, Ho FK, Reynolds L, Lau KK, Crystal S, Lau WCY, Man KKC, Brauer R, Chan EW, Shen CY, Kim JH, Lum TYS, Hartikainen S, Koponen M, Rooke E, Bazelier M, Klungel O, Setoguchi S, Pell JP, Cook S and Wong ICK (2020) Application of healthcare 'Big data' in CNS drug research: the example of the neurological and mental health global epidemiology network (NeuroGEN). CNS Drugs 34, 897–913.
- Lai EC, Man KK, Chaiyakunapruk N, Cheng CL, Chien HC, Chui CS, Dilokthornsakul P, Hardy NC, Hsieh CY, Hsu CY, Kubota K, Lin TC, Liu Y, Park BJ, Pratt N, Roughead EE, Shin JY, Watcharathanakij S, Wen J, Wong IC, Yang YH, Zhang Y and Setoguchi S (2015) Brief report: databases in the Asia-pacific region: the potential for a distributed network approach. *Epidemiology* 26, 815–820.
- Lai EC, Ryan P, Zhang Y, Schuemie M, Hardy NC, Kamijima Y, Kimura S, Kubota K, Man KK, Cho SY, Park RW, Stang P, Su CC, Wong IC, Kao YY and Setoguchi S (2018a) Applying a common data model to Asian databases for multinational pharmacoepidemiologic studies: opportunities and challenges. *Clinical Epidemiology* 10, 875–885.
- Lai EC, Shin JY, Kubota K, Man KKC, Park BJ, Pratt N, Roughead EE, Wong ICK, Kao Yang YH and Setoguchi S (2018b) Comparative safety of NSAIDs for gastrointestinal events in Asia-pacific populations: a multidatabase, international cohort study. *Pharmacoepidemiology and Drug Safety* 27, 1223–1230.
- Lao KSJ, Tam AWY, Wong ICK, Besag FMC, Man KKC, Chui CSL and Chan EW (2017) Prescribing trends and indications of antipsychotic medication in Hong Kong from 2004 to 2014: general and vulnerable patient groups. *Pharmacoepidemiology and Drug Safety* 26, 1387–1394.
- Lee H, Song DH, Han E and Kang HY (2018) Nationwide epidemiologic study of atypical antipsychotic use Among pediatric population with mental

illness in Korea. J Journal of Child and Adolescent Psychopharmacology 28, 205–215.

- Ma X, Maimaitirexiati T, Zhang R, Gui X, Zhang W, Xu G and Hu G (2014) HTR2C Polymorphisms, olanzapine-induced weight gain and antipsychoticinduced metabolic syndrome in schizophrenia patients: a meta-analysis. *International Journal of Psychiatry in Clinical Practice* 18, 229–242.
- Man KKC, Shao SC, Chaiyakunapruk N, Dilokthornsakul P, Kubota K, Li J, Ooba N, Pratt N, Pottegård A, Rasmussen L, Roughead EE, Shin JY, Su CC, Wong ICK, Kao Yang YH and Lai EC (2020) Metabolic events associated with the use of antipsychotics in children, adolescents and young adults: a multinational sequence symmetry study. *European Child and Adolescent Psychiatry*. doi: 10.1007/s00787-020-01674-6. doi: Epub ahead of print.
- McIntyre RS and Jerrell JM (2008) Metabolic and cardiovascular adverse events associated with antipsychotic treatment in children and adolescents. *Archives of Pediatrics & Adolescent Medicine* 162, 929–935.
- Mulder H, Cohen D, Scheffer H, Gispen-de Wied C, Arends J, Wilmink FW, Franke B and Egberts AC (2009) HTR2C Gene polymorphisms and the metabolic syndrome in patients with schizophrenia: a replication study. *Journal of Clinical Psychopharmacology* **29**, 16–20.
- Olfson M, King M and Schoenbaum M (2015) Treatment of young people with antipsychotic medications in the United States. JAMA Psychiatry 72, 867–874.

- Petersen I, Douglas I and Whitaker H (2016) Self controlled case series methods: an alternative to standard epidemiological study designs. *British Medical Journal* 354, i4515.
- Pillay J, Boylan K, Newton A, Hartling L, Vandermeer B, Nuspl M, MacGregor T, Featherstone R and Carrey N (2018) Harms of antipsychotics in children and young adults: a systematic review update. *Canadian Journal of Psychiatry* 63, 661–678.
- Pillinger T, McCutcheon RA, Vano L, Mizuno Y, Arumuham A, Hindley G, Beck K, Natesan S, Efthimiou O, Cipriani A and Howes OD (2020) Comparative effects of 18 antipsychotics on metabolic function in patients with schizophrenia, predictors of metabolic dysregulation, and association with psychopathology: a systematic review and network meta-analysis. *The Lancet. Psychiatry* 7, 64–77.
- Sjo CP, Stenstrøm AD, Bojesen AB, Frølich JS and Bilenberg N (2017) Development of metabolic syndrome in drug-naive adolescents after 12 months of second-generation antipsychotic treatment. *Journal of Child and Adolescent Psychopharmacology* 27, 884–891.
- Vandenberghe F, Najar-Giroud A, Holzer L, Conus P, Eap CB and Ambresin AE (2018) Second-generation antipsychotics in adolescent psychiatric patients: metabolic effects and impact of an early weight change to predict longer term weight gain. *Journal of Child and Adolescent Psychopharmacology* 28, 258–265.