ON PARRONDO'S PARADOX: HOW TO CONSTRUCT UNFAIR GAMES BY COMPOSING FAIR GAMES

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Abstract

We construct games of chance from simpler games of chance. We show that it may happen that the simpler games of chance are fair or unfavourable to a player and yet the new combined game is favourable—this is a counter-intuitive phenomenon known as Parrondo's paradox. We observe that all of the games in question are random walks in periodic environments (RWPE) when viewed on the proper time scale. Consequently, we use RWPE techniques to derive conditions under which Parrondo's paradox occurs.

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1. Introduction

Parrondian strategies are where losing games can cooperate to win [6]. The original example of Parrondo's games consist of two coin tossing games. Game A consists of Coin 1 biased to lose. Game B consists of two coins—Coin 2 with losing bias and Coin 3 with winning bias—but a state-dependent rule is chosen to favour the losing Coin 2. Hence both games A and B are losing games. However when A and B are alternated in a deterministic or even random manner, surprisingly the player has a winning expectation. This effect has been interpreted in terms of a

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discrete-time Brownian ratchet at length elsewhere [7], where conventional Brownian ratchets [2] have been the inspiration. An alternative view, which we call the Boston interpretation [15], recognises that although game B favours Coin 2 with losing bias, if the state-dependence is removed, game B now favours the winning Coin 3—then when games A and B are mixed, game A has the affect of randomisation or 'break-up' of game B's state-dependence, thus tilting favour towards Coin 3 with winning bias. This explanation was also independently deduced by J. Maynard Smith [19]. In addition to the ratchet interpretation and the Boston interpretation, this paper will examine another viewpoint by considering the process as a random walk in a periodic environment (RWPE).

We now briefly summarise the literature on Parrondo's games. In [9] the state-dependent rule for game B is to choose Coin 2 if the player's capital is a multiple of some positive integer M—analysis showed the paradox could be observed for general values of M. In [22] it is shown that the paradox can be observed when both games A and B are multiple coin games. In [23] we have the first analysis of Parrondo's games in terms of Shannon entropy and in [10] the entropy parameter spaces are graphically displayed. The probability parameter space is shown in [8]. In [15] a minimal three-state game with asymmetric transition probabilities is analysed and in [21] state-dependence on capital is replaced by dependence on the past history of the game, leading to a larger probability parameter space.

The surge of interest in analysing Parrondian games is motivated by a number of areas. Information theorists have long studied the problem of producing a fair game from biased coins [4] and the roots of this can be traced back to the work of [29]—Parrondo's games go a step further in producing a winning game from losing games. Siegman [28] has reinterpreted 'capital' of the games in terms of electron occupancies of energy levels—the paradox can then be reproduced using the rate equation approach typically used in laser analysis. In the physical world there are many types of processes where losing helps to win, such as a sacrifice in the game of chess or a valley in the fitness landscape of an animal species. Many biological effects are linked to ratchet-type phenomena and [30] have analysed enzyme transport with a four-state model. Applicability to population genetics, evolution and economics has been suggested [18]. In finance, [17] have shown that under certain conditions capital can grow by investing in an asset with *negative* typical growth rate. Quantum ratchets have now been experimentally realised [16] and recasting Parrondian games, based on ratchet phenomena, as quantum games [3, 5, 20] is thus of interest.

In control theory, it can be shown that the combination of two unstable systems can become stable [1]. Velocity of propagation through an array of coupled oscillators, under certain conditions, can *increase* even though the damping coefficient is *increased* [27]. In the area of granular flow, drift can occur in a counter-intuitive direction such as is exemplified in the famous Brazil nut paradox [26]. Also declining

branching processes can be combined to increase [12]. In [25] it is shown that noisy information can sometimes be better than clean information. In [24] it is shown that with switched diffusion processes in random media it is possible to get a positive-recurrent processes (that is, with no drift) from mixed transient processes (that is, with drifts all in the same direction)—this is almost certainly a continuous-time analogue to the Parrondian discrete-time process. Assuming we construct Parrondo's games to only deal in transactions of one unit of capital per event, then we have a skip-free process, and a statistical interpretation of the central result is that declining birth-death processes can be combined to form an increase.

In this paper we further investigate Parrondo's paradox. We construct a class of composite games and investigate their fairness by formulating the problem in the language of random walks in periodic environments (RWPE). We find many new, interesting and counter-intuitive results.

2. Mathematical tools

We construct a composite game from two simple games A and B which are modeled as RWPEs. These two games can be combined in two ways: deterministically or stochastically.

If we have played n times, and n is divisible by k (for an integer k), we then play game A. If we have played n times and n is not divisible by k, we play game B. Thus games A and B are alternated in a deterministic pattern. We denote by Y_n our capital after n plays of this game.

To alternate games A and B randomly we toss a coin with probability p of heads. If the coin comes up heads, we play game A, and if tails, we play game B. We denote by Z_n our capital after n repetitions.

The sequence of values of our capital in either of these games is a Markov random walk which changes by ± 1 in each epoch. The two random walks Z_n and Y_n differ in that Z_n is time homogeneous, while Y_n is not time homogeneous. Moreover, Z_n is a random walk in a periodic environment. The process Y_n is not a RWPE, but the process $Y_n' \equiv Y_{kn}$ is a RWPE.

As shown in Section 4 we extend this construction to define composite games from more than two simple games. In each case we identify the capital of the player as a RWPE.

We say that a game is fair, winning or losing if the random walk for the capital of a player, X_n , is recurrent, or transient to ∞ or to $-\infty$, respectively. That is, a Markov chain X_n is

recurrent (fair) if
$$P\left\{-\infty = \lim \inf_{n \to \infty} X_n < \lim \sup_{n \to \infty} X_n = \infty\right\} = 1;$$

[4]

transient to
$$\infty$$
 (winning) if $P\left\{\lim_{n\to\infty}X_n=+\infty\right\}=1$;
transient to $-\infty$ (losing) if $P\left\{\lim_{n\to\infty}X_n=-\infty\right\}=1$.

We note that the characterisation of a game as fair, winning, or losing by the traditional comparisons

$$E[X_{n+1} \mid X_n] = X_n$$
, $E[X_{n+1} \mid X_n] > X_n$ and $E[X_{n+1} \mid X_n] < X_n$,

respectively, does not cover the behaviour of all random walks, in particular RWPE's.

2.1. Key's criterion We consider a time homogeneous random walk, X_n , in an N-periodic environment, or equivalently we have a state dependent random game W. We assume that the maximal step size in the positive direction is R, while in the negative direction it is L, that is, $P\{X_{n+1} \in \{-L+k,\ldots,x+R\} \mid X_n=k\}=1$. Moreover, for each k the maximum right and left step sizes are always possible, that is, $P\{X_{n+1} = -L+k \mid X_n=k\}P\{X_{n+1} = R+k \mid X_n=k\} > 0$. Given the environment, the walk X_n obeys the backward master equation

$$P\{\cdot \mid X_n = k\} = \sum_{j=-L}^R e(k,j) P\{\cdot \mid X_{n+1} = j+k\}, \tag{2.1}$$

where $e(k, j) \equiv P\{X_{n+1} = j + k | X_n = k\}$ denotes the transition probability from the state k to j + k in one time epoch. We denote by $f_{k-i} \equiv P\{\cdot \mid X_{n+1} = k - i\}$ (since X_n is time homogeneous f does not depend on n) and rewrite (2.1) as a system for the vector $[f_{-L+k}, f_{-L+1+k}, \dots, f_{R-1+k}]^T$ with the matrix A_k whose entries are given by

$$\mathbf{A}_{k}[i, j] = \begin{cases} -e(k, -L + j)/e(k, -L) & \text{if } i = 1, \ j \neq L, \\ (1 - e(k, 0))/e(k, -L) & \text{if } i = 1, \ j = L, \\ 1 & \text{if } i \geq 2, \ j = i - 1, \\ 0 & \text{otherwise.} \end{cases}$$

We also define the matrix $M = A_1 A_2 \cdots A_N$.

According to [11] we define constants d_i , $i=1,2,\ldots,R+L$ as follows. For each eigenvalue λ_i of **M** (including multiplicities), we put $d_i = \log(|\lambda_i|)$, and we list the d_i 's in increasing order, so that $d_1 \le d_2 \le \cdots \le d_{R+L}$. Then

- if $ln(c(\mathbf{W})) \equiv d_R + d_{R+1} > 0$ then the RWPE X_n is transient to ∞ ;
- if $ln(c(\mathbf{W})) \equiv d_R + d_{R+1} = 0$ then the RWPE X_n is recurrent;

if
$$\ln(c(\mathbf{W})) \equiv d_R + d_{R+1} < 0$$
 then the RWPE X_n is transient to $-\infty$. (2.2)

It is shown in [13] that X_n is recurrent if the characteristic polynomial of M has a double root at 1.

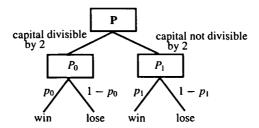


FIGURE 1. Schematic representation of a 2-(state) periodic game $P = (P_0, P_1)$. Games composed of such games are discussed in Section 3.

3. Games composed of 2-periodic games

First we consider an example of two 2-periodic games $P = (P_0, P_1)$ and $Q = (Q_0, Q_1)$; that is, we have two RWPE's, each of them 2-periodic (in space). If X_n denotes the capital of a player at time n who plays the game according to the rule P, then the transition probabilities are given by

$$P\{X_{n+1} = x + 1 \mid X_n = x\} = \begin{cases} p_0 & \text{if } x = 0 \mod 2, \\ p_1 & \text{if } x = 1 \mod 2, \end{cases}$$

$$P\{X_{n+1} = x - 1 \mid X_n = x\} = \begin{cases} 1 - p_0 & \text{if } x = 0 \mod 2, \\ 1 - p_1 & \text{if } x = 1 \mod 2; \end{cases}$$
(3.1)

that is, if the capital is even it changes according to P_0 , and if it is odd it changes according to P_1 , see, for example, Figure 1. The transition probabilities for the RWPE governed by \mathbf{Q} are given by formulas analogous to (3.1), with p's replaced by q's, with the rule Q_0 and Q_1 at the even and odd positions, respectively.

To employ the criterion of [11] to the RWPE governed by $\bf P$ we construct the matrix $\bf M$

$$\mathbf{M} = \mathbf{A}_1 \mathbf{A}_0 \equiv \begin{bmatrix} 1/(1-p_1) & -p_1/(1-p_1) \\ 1 & 0 \end{bmatrix} \begin{bmatrix} 1/(1-p_0) & -p_0/(1-p_0) \\ 1 & 0 \end{bmatrix}.$$

Since here R = L = 1, $d_R + d_{R+1} = d_1 + d_2 = \log(|\det(\mathbf{M})|)$, so we find that the RWPE governed by **P** is transient to $-\infty$, recurrent, or transient to $+\infty$ if

$$c(\mathbf{P}) \equiv c(P_0, P_1) \equiv \frac{p_0 p_1}{(1 - p_0)(1 - p_1)} < (=)(>)1.$$

We note that this analysis easily extends to any period N of the environment, see, for example, Section 4.

3.1. An example of a pseudo-paradox We suppose that the games P and Q are fair, that is,

$$c(P_0, P_1) = \frac{p_0(1 - p_0)}{p_1(1 - p_1)} = 1$$
 and $c(Q_0, Q_1) = \frac{q_0(1 - q_0)}{q_1(1 - q_1)} = 1$.

If we alternate (deterministically) games P and Q in that very order, then, starting at the origin, the composite game may be fair (losing) (winning) if

$$c(\mathbf{P}, \mathbf{Q}) = p_0 q_1 / ((1 - p_0)(1 - q_1)) = (<)(>)1.$$

Under this strategy we play the game P_0 when the winnings are even and the game Q_1 when the winnings are odd. The games P_1 and Q_0 are never played. Hence $c(\mathbf{P}, \mathbf{Q}) = c(P_0, Q_1)$, and the fact that the composite game is losing when it is constructed from two winning games is just an apparent paradox. If we play the same two games when starting at an odd position then $c(\mathbf{P}, \mathbf{Q}) = c(P_1, Q_0) = 1/c(P_0, Q_1)$, and we have a winning game constructed out of two winning games.

3.2. Effects of randomisation Now we construct a RWPE, X_n , by choosing at random between **P** and **Q**. If our capital is even then a coin with probability of heads g_0 is used to choose whether the game P_0 or Q_0 will be played; if our capital is odd then a coin with probability of heads g_1 is used to choose between the games P_1 and Q_1 . In this scheme all four games are played. The transition probabilities are given by

$$P\{X_{n+1} = x + 1 \mid X_n = x\} = \begin{cases} g_0 p_0 + (1 - g_0) q_0 & \text{if } x = 0 \text{ mod } 2, \\ g_1 p_1 + (1 - g_1) q_1 & \text{if } x = 1 \text{ mod } 2. \end{cases}$$
(3.2)

The fairness of the composite game is determined by the factor $c(\mathbf{P}, \mathbf{Q})$ given by

$$c(\mathbf{P}, \mathbf{Q}) = \frac{g_0 p_0 + (1 - g_0) q_0}{(1 - g_0 p_0 - (1 - g_0) q_0)} \frac{g_1 p_1 + (1 - g_1) q_1}{(1 - g_1 p_1 - (1 - g_1) q_1)}.$$
 (3.3)

By direct simplification of (3.3) we observe that if $g_0 = g_1$ (and the games **P** and **Q** are fair) then the composite game is fair. Also, if $p_1 = q_1$ then the composite game is fair. However, for other values of parameters two fair games can be used to compose a game which is winning, losing or fair depending on how the two simple games are randomised. Figure 2 illustrates this statement. That is, randomisation can produce Parrondo's paradox.

We also consider a game composed from two unfair games P and Q. Games P and Q are taken to be unfair the same way, that is, both are losing or both are winning. We note that when one coin is used to choose whether P or Q is played, (that is, when $g_0 = g_1$ in (3.2)) then the composed game is always unfair. However, if two coins are used then the randomised game may be winning when both simple games P and Q are losing. This example of Parrondo's paradox is illustrated in Figure 3.

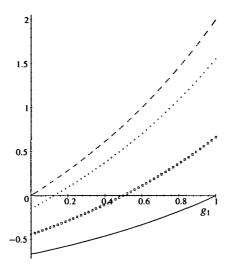


FIGURE 2. We take two 2-periodic fair games with $p_1 = 1/2$ and $q_1 = 1/4$, and plot $c(\mathbf{P}, \mathbf{Q}) - 1$ in (3.3) as a function of g_1 for various values of g_0 ; that is, the composite game is winning for positive values on the graph, and losing for negative values. We have: dashes: $g_0 = 0$; crosses: $g_0 = 1/8$; boxes: $g_0 = 1/2$; solid: $g_0 = 1$.

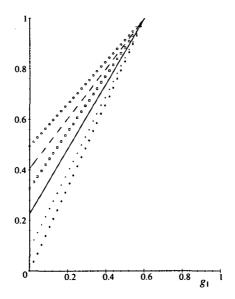


FIGURE 3. We take two 2-periodic losing games with $p_0 = 0.675$, $p_1 = 0.1$, $q_1 = 0.75$ and plot the fairness curve $c(\mathbf{P}, \mathbf{Q}) = 1$ in (3.3) in the (g_1, g_0) plane for various values of q_0 : circles: $q_0 = 0$; dashes: $q_0 = 0.075$; boxes: $q_0 = 0.125$; solid: $q_0 = 0.175$; crosses: $q_0 = 0.225$; diamonds: $q_0 = 0.25$. The composite game is winning when g_0 and g_1 are selected above the fairness curve.

3.3. Three 2-periodic games We consider three 2-periodic games P, Q and R, each defined analogously to (3.1), with transition probabilities given in terms of p_0 , p_1 , q_0 , q_1 , and r_0 , r_1 , respectively. We construct the composite game PQR. We investigate whether the composite game may be winning (losing) if the individual games P, Q and R are fair. If the capital of the player at time n is X_n then the capital at time 3n, $Y_n \equiv X_{3n}$, is a random walk in a 2-periodic environment, taking steps of ± 3 , ± 1 . Specifically, we have

$$P\{Y_{n+1} = x + 3 \mid Y_n = x\} = \begin{cases} a_0 \equiv p_0 q_1 r_0 & \text{if } x = 0 \text{ mod } 2, \\ a_1 \equiv p_1 q_0 r_1 & \text{if } x = 1 \text{ mod } 2, \end{cases}$$

$$P\{Y_{n+1} = x - 3 \mid Y_n = x\} = \begin{cases} b_0 \equiv (1 - p_0)(1 - q_1)(1 - r_0) & \text{if } x = 0 \text{ mod } 2, \\ b_1 \equiv (1 - p_1)(1 - q_0)(1 - r_1) & \text{if } x = 1 \text{ mod } 2, \end{cases}$$

$$P\{Y_{n+1} = x + 1 \mid Y_n = x\} = \begin{cases} c_0 \equiv p_0 q_1(1 - r_0) + p_0(1 - q_1) r_0 \\ + (1 - p_0) q_1 r_0 & \text{if } x = 0 \text{ mod } 2, \end{cases}$$

$$c_1 \equiv p_1 q_0(1 - r_1) + p_1(1 - q_0) r_1 \\ + (1 - p_1) q_0 r_1 & \text{if } x = 1 \text{ mod } 2, \end{cases}$$

$$P\{Y_{n+1} = x - 1 \mid Y_n = x\} = \begin{cases} d_0 \equiv (1 - p_0)(1 - q_1) r_0 + (1 - p_0) q_1(1 - r_0) \\ + p_0(1 - q_1)(1 - r_0) & \text{if } x = 0 \text{ mod } 2, \end{cases}$$

$$d_1 \equiv (1 - p_1)(1 - q_0) r_1 + (1 - p_1) q_0(1 - r_1) \\ + p_1(1 - q_0)(1 - r_1) & \text{if } x = 1 \text{ mod } 2. \end{cases}$$

Moreover, the random walk $Z_n \equiv Y_{2n}$ is an ordinary random walk, that is, a sum of i.i.d. random variables taking values ± 6 , ± 4 , ± 2 , and 0, and the transition probabilities are

$$\begin{split} P\{Z_{n+1} &= z + 6 \mid Z_n = z\} = a_0 a_1, \\ P\{Z_{n+1} &= z - 6 \mid Z_n = z\} = b_0 b_1, \\ P\{Z_{n+1} &= z + 4 \mid Z_n = z\} = a_0 c_1 + c_0 a_1, \\ P\{Z_{n+1} &= z + 4 \mid Z_n = z\} = b_0 d_1 + d_0 b_1, \\ P\{Z_{n+1} &= z + 2 \mid Z_n = z\} = c_0 c_1 + a_0 d_1 + a_1 d_0, \\ P\{Z_{n+1} &= z - 2 \mid Z_n = z\} = d_0 d_1 + b_0 c_1 + b_1 c_0, \\ P\{Z_{n+1} &= z \mid Z_n = z\} = d_0 c_1 + d_1 c_0 + a_0 b_1 + a_1 b_0. \end{split}$$

If the games P, Q and R are fair, then by direct calculations, $EZ_n = 0$, so the composite game is fair, and there is no paradox in this case.

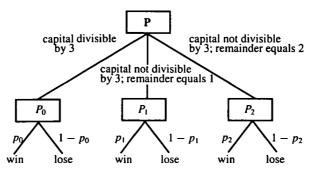


FIGURE 4. Schematic representation of a 3-(state) periodic game $P = (P_0, P_1, P_2)$. Games composed of such games are discussed in Section 4.

4. Games composed of 3-periodic games

We consider games composed of 3-periodic games P, Q, and R. We define a 3-periodic game $P = (P_0, P_1, P_2)$ by its transition probabilities as

$$P\{X_{n+1} = x + 1 \mid X_n = x\} = \begin{cases} p_0 & \text{if } x = 0 \text{ mod } 3, \\ p_1 & \text{if } x = 1 \text{ mod } 3, \\ p_2 & \text{if } x = 2 \text{ mod } 3, \end{cases}$$
(4.1)

compare Figure 4. Transition probabilities of the games \mathbf{Q} and \mathbf{R} are defined by formulas analogous to (4.1) with p's replaced by q's and r's, respectively. According to [11], to determine the fairness of the game \mathbf{P} we consider the product of the eigenvalues of the matrix $\mathbf{M} \equiv \mathbf{A}_1 \mathbf{A}_2 \mathbf{A}_0$ where

$$\mathbf{A}_i = \begin{bmatrix} 1/(1-p_i) & -p_i/(1-p_i) \\ 1 & 0 \end{bmatrix} \quad i = 0, 1, 2.$$

If $p_0 p_1 p_2/[(1-p_0)(1-p_1)(1-p_2)] = (<)(>)1$ then the game **P** is fair (losing) (winning).

4.1. Two 3-periodic games To analyse a game (P, Q), where P and Q are 3-periodic games, we construct a random walk $Y_n \equiv X_{2n}$ where X_n denotes the position of the walker at time n. The transition probabilities for Y_n are given by

$$P\{Y_{n+1} = x + 2 \mid Y_n = x\} = \begin{cases} a_0 \equiv p_0 q_1 & \text{if } x = 0 \mod 3, \\ a_1 \equiv p_1 q_2 & \text{if } x = 1 \mod 3, \\ a_2 \equiv p_2 q_0 & \text{if } x = 2 \mod 3, \end{cases}$$

$$P\{Y_{n+1} = x \mid Y_n = x\} = \begin{cases} b_0 \equiv p_0(1 - q_1) + (1 - p_0)q_2 & \text{if } x = 0 \bmod 3, \\ b_1 \equiv p_1(1 - q_2) + (1 - p_1)q_0 & \text{if } x = 1 \bmod 3, \\ b_2 \equiv p_2(1 - q_0) + (1 - p_2)q_1 & \text{if } x = 2 \bmod 3, \end{cases}$$

$$P\{Y_{n+1} = x - 2 \mid Y_n = x\} = \begin{cases} c_0 \equiv (1 - p_0)(1 - q_2) & \text{if } x = 0 \bmod 3, \\ c_1 \equiv (1 - p_1(1 - q_0)) & \text{if } x = 1 \bmod 3, \\ c_2 \equiv (1 - p_2)(1 - q_1) & \text{if } x = 2 \bmod 3. \end{cases}$$

Hence Y_n is a 3-periodic random walk taking steps ± 2 and zero. To employ Key's criterion most efficiently we observe that it is sufficient to analyse Y_n as a RWPE on the even integers which visits only nearest neighbours. To this end we construct the 2×2 matrix $\mathbf{M} = \mathbf{A}_1 \mathbf{A}_2 \mathbf{A}_0$ where

$$\mathbf{A}_i = \begin{bmatrix} (1-b_i)/c_i & -a_i/c_i \\ 1 & 0 \end{bmatrix} \quad i = 0, 1, 2.$$

The determinant of M is given by

$$\det(\mathbf{M}) = \frac{p_0 p_1 p_2 q_0 q_1 q_2}{(1 - p_0)(1 - p_1)(1 - p_2)(1 - q_0)(1 - q_1)(1 - q_2)}.$$

Hence if the games P and Q are fair, the composite game is fair and no paradox is observed.

We note that this reduction in the dimension of M occurs whenever the temporal period is even, since the step sizes of the derived process Y_n are even. Specifically, if the temporal period is T is even then M can be taken to be the product of $T \times T$ matrices, and when the temporal period T is odd, then M is the product of $2T \times 2T$ matrices.

4.2. Two 3-periodic games randomised We construct a composite game from two 3-periodic random games $P = (P_0, P_1, P_2)$ and $Q = (Q_0, Q_1, Q_2)$ by selecting at random the game to be played at each step, with probability of playing P_i equal to g_i , i = 0, 1, 2. That is, we define a 3-periodic random walk Y_n taking values ± 1 with transition probabilities

$$P\{Y_{n+1} = x + 1 \mid Y_n = x\} = \begin{cases} \rho_0 \equiv g_0 p_0 + (1 - g_0) q_0 & \text{if } x = 0 \text{ mod } 3, \\ \rho_1 \equiv g_1 p_1 + (1 - g_1) q_1 & \text{if } x = 1 \text{ mod } 3, \\ \rho_2 \equiv g_2 p_2 + (1 - g_2) q_2 & \text{if } x = 2 \text{ mod } 3. \end{cases}$$
(4.2)

The walk Y_n is recurrent (transient to ∞) (transient to $-\infty$) according to

$$c(\mathbf{P}, \mathbf{Q}) \equiv \frac{\rho_0 \rho_1 \rho_2}{(1 - \rho_0)(1 - \rho_1)(1 - \rho_2)} = (>)(<)1.$$

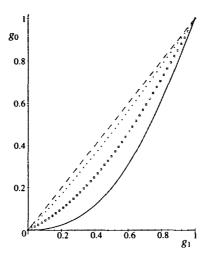


FIGURE 5. We take two fair 3-periodic games with $g_1 = g_2$, $p_1 = p_2 = 1/2$, and $q_1 = q_2$. For a fixed value of q_1 we plot curves along which the composite game is fair, in the space (g_1, g_0) . As q_1 varies from 0 to 1/2 the composite game is losing above and winning below its corresponding "fairness" curve; solid: $q_1 = 0$; boxes: $q_1 = 1/6$; crosses: $q_1 = 1/3$; dashes: the line $g_0 = g_1$. As q_1 varies from 1/2 to 1 the composite game is winning below and losing above its corresponding "fairness" curve. In this case the problem is symmetric, so that we have: solid: $q_1 = 1$; boxes $q_1 = 5/6$; crosses: $q_1 = 2/3$.

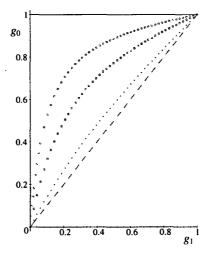


FIGURE 6. We take two fair 3-periodic games with $g_1 = g_2$, $p_1 = p_2 = 1/2$, and $q_0 = q_2$. We plot the curves along which the composite game is fair in the space (g_1, g_0) , for a fixed value of q_0 . As q_0 varies from 0 to 1/2, the game is winning above and losing below each of its "fairness" curves; solid: $q_0 = 0$; circles: $q_0 = 1/12$; boxes: $q_0 = 1/6$; crosses: $q_0 = 1/3$; dashes: the line $g_0 = g_1$. As q_0 increases from 1/2 to 1 the game is losing above and winning below its fairness curve. The case of this example is symmetric so that we have: crosses: $q_0 = 2/3$; boxes: $q_0 = 5/6$; circles: $q_0 = 11/12$.

We investigate the value of $c(\mathbf{P}, \mathbf{Q})$ for various values of the parameters. We note that if the games \mathbf{P} and \mathbf{Q} are fair with $p_1 = p_2 = q_1 = q_2$ then the randomised composite game is fair for all possible choices of g_i 's. Next we take two fair games with fixed values $p_1 = p_2$ and $q_1 = q_2$, that is, $\mathbf{P} = (P_0, P_1, P_1)$ and $\mathbf{Q} = (Q_0, Q_1, Q_1)$.

We investigate how randomisation affects the fairness of the composite game. As shown in Figure 5 there is a wide range of parameters for which the randomised game is unfair, so that randomisation induces the paradox. We note that the range of the randomisation parameters g's is affected by the way games P and Q are structured. In the examples illustrated in Figure 6 we take two fair games $P = (P_0, P_1, P_1)$ and $Q = (Q_0, Q_1, Q_0)$ and we note that the range of parameters $g_0, g_1 = g_2$ for which the randomised game composed from these "shifted" games is larger compared to the previous "not-shifted" case, see, for example, Figures 5 and 6.

4.3. Three 3-periodic games First we construct a game (PQR) in which the 3-periodic games P, Q and R are played in a deterministic order. We show that if all three games are fair then the resulting game may be losing. We also show that if all three games are winning then the composite game may be losing. If X_n denotes the winnings of the composite game at time n then its transition probabilities are given by:

if $x = 0 \mod 3$ then

$$\begin{split} P\{X_{n+3} = x + 3 \mid X_n = x\} &= a_0 \equiv p_0 q_1 r_2, \\ P\{X_{n+3} = x - 3 \mid X_n = x\} &= b_0 \equiv (1 - p_0)(1 - q_2)(1 - r_1), \\ P\{X_{n+3} = x + 1 \mid X_n = x\} &= c_0 \equiv p_0 (1 - q_2) r_1 \\ &\qquad \qquad + (1 - p_0) q_2 (1 - r_0) + p_0 (1 - q_1)(1 - r_0), \\ P\{X_{n+3} = x - 1 \mid X_n = x\} &= d_0 \equiv (1 - p_0)(1 - q_2) r_1 \\ &\qquad \qquad + (1 - p_0) q_2 (1 - r_0) + p_0 (1 - q_1)(1 - r_0); \end{split}$$

if $x = 1 \mod 3$ then

$$P\{X_{n+3} = x + 3 \mid X_n = x\} = a_1 \equiv p_1 q_2 r_0,$$

$$P\{X_{n+3} = x - 3 \mid X_n = x\} = b_1 \equiv (1 - p_1)(1 - q_0)(1 - r_2),$$

$$P\{X_{n+3} = x + 1 \mid X_n = x\} = c_1 \equiv p_1 q_2 (1 - r_0) + p_1 (1 - q_2) r_1 + (1 - p_1) q_0 r_1,$$

$$P\{X_{n+3} = x - 1 \mid X_n = x\} = d_1 \equiv (1 - p_1)(1 - q_0) r_2 + (1 - p_1) q_0 (1 - r_1) + p_1 (1 - q_2)(1 - r_1);$$

if $x = 2 \mod 3$ then

$$P\{X_{n+3} = x + 3 \mid X_n = x\} = p_2 q_0 r_1 \equiv a_2,$$

$$P\{X_{n+3} = x - 3 \mid X_n = x\} = (1 - p_2)(1 - q_1)(1 - r_0) \equiv b_2,$$

$$P\{X_{n+3} = x + 1 \mid X_n = x\} = p_2 q_0 (1 - r_1) + p_2 (1 - q_0) r_2 + (1 - p_2) q_1 r_2 \equiv c_2,$$

$$P\{X_{n+3} = x - 1 \mid X_n = x\} = d_2 \equiv (1 - p_2)(1 - q_1) r_0 + (1 - p_2) q_1 (1 - r_2) + p_2 (1 - q_0)(1 - r_2).$$

Hence we define the walk in a periodic environment by $Y_n \equiv X_{3n}$, and determine its transient/recurrent properties by investigating locations of the eigenvalues of the matrix $\mathbf{M} = \mathbf{A}_1 \mathbf{A}_2 \mathbf{A}_0$. Since the walk Y_n takes on the values ± 3 , ± 1 , R = L = 3, each of the matrices \mathbf{A}_i , i = 1, 2, 3, is a 6×6 matrix given by

$$\mathbf{A}_{i} = \begin{bmatrix} 0 & -d_{i}/b_{i} & 1/b_{i} & -c_{i}/b_{i} & 0 & -a_{i}/b_{i} \\ 1 & 0 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 0 & 1 & 0 \end{bmatrix}, \quad i = 0, 1, 2.$$

The examples illustrated in Figure 7 and 8 show that the game composed deterministically from fair (non-symmetric) games can be fair or losing or winning, depending on parameters of the problem. This form of the paradox is not observed when composing (deterministically) two or three 2-periodic games. We observe even more interesting behaviour when composing deterministically unfair (losing) games. Figures 9 and 10 show such examples.

By varying just one parameter, it is possible to change the composed game from losing to winning, change it back to losing, and then back again to winning, see, for example, Figure 9. Three state transitions of the composed game have not been observed before.

5. Conclusions

Given the highly nonlinear nature of the recurrence/transience criteria for random walks in periodic environments as a function of the process parameters, it is not surprising that any scheme that combines recurrent random walks in periodic environment to produce new random walks in a periodic environments will produce transient processes, except in trivial cases.

What is surprising to us is that one can find non-trivial cases where random walks in periodic environments with negative drift can be combined to form random walks in periodic environments with positive drift, Parrondo's paradox. We believe that these reversals arise because combining these processes in the manner described above radically changes the frequency at which transitions are governed by probability distributions with positive means, and that future investigations should center on investigating these frequencies. It is likely that it will be easier to understand this phenomenon in the case of stochastic combinations of the processes since these always lead to the study of nearest neighbour random walks in periodic environments.

When a composite game is constructed from simple 3-periodic games, Parrondo's paradox is much more interesting and complex than in the case of 2-periodic games. As

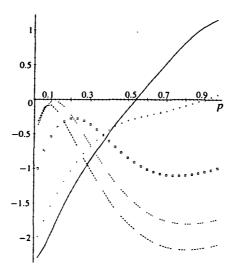


FIGURE 7. We take 3 fair games with $p_0 = p_1 \equiv p$, $q_0 = q_1 \equiv q$, and $r_0 = r_1 \equiv r$, and plot $\ln(c(\mathbf{P}, \mathbf{Q}, \mathbf{R})) = d_3 + d_4$, compare (2.2), as a function of p, for q = 0.1 and various values of r, solid: r = 3/4; crosses: r = 1/2, boxes: r = 1/4; short dashes: r = 1/8; long dashes: r = 1/16. We note that the composite game is always losing, no matter what p is, for a wide range of values of the parameter r. If $\ln(c(\mathbf{P}, \mathbf{Q}, \mathbf{R}))$ is positive (zero) (negative) then the game $(\mathbf{P}, \mathbf{Q}, \mathbf{R})$) is winning (fair) (losing).

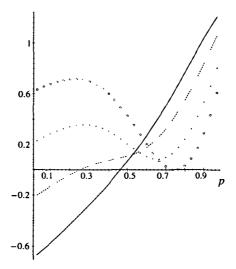


FIGURE 8. We take 3 fair games with $p_0 = p_1 \equiv p$, $q_0 = q_1 \equiv q$, and $r_0 = r_1 \equiv r$, and plot $\ln(c(\mathbf{P}, \mathbf{Q}, \mathbf{R})) = d_3 + d_4$, compare (2.2), as a function of p, for q = 3/4 and various values of r: circles: r = 3/4, crosses: r = 5/8; dashes: r = 1/2; solid: r = 3/8. If $\ln(c(\mathbf{P}, \mathbf{Q}, \mathbf{R}))$ is positive (zero) (negative) then the game $(\mathbf{P}, \mathbf{Q}, \mathbf{R})$) is winning (fair) (losing).

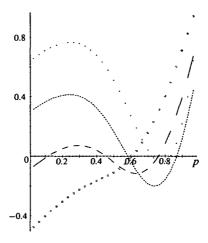


FIGURE 9. We take 3 losing games with $p_0 = p_1 \equiv p$, $q_0 = q_1 \equiv q$, and $r_0 = r_1 \equiv r$, and $p_2 = 0.8(1-p)^2/((1-p)^2+p^2)$, $q_2 = 0.8(1-q)^2/((1-q)^2+q^2)$, and $r_2 = 0.9(1-r)^2/((1-r)^2+r^2)$. We set q = 3/4, so that $q_2 = 0.08$, and plot $\ln(c(\mathbf{P}, \mathbf{Q}, \mathbf{R})) = d_3 + d_4$, compare (2.2), as a function of p for various values of r, circles: r = 1/2, $(r_2 = 0.45)$; solid: r = 5/8, $(r_2 = 0.238235...)$; dashes: r = 3/4, $(r_2 = 0.09)$; crosses r = 7/8, $(r_2 = 0.018)$. If $\ln(c(\mathbf{P}, \mathbf{Q}, \mathbf{R}))$ is positive (zero) (negative) then the game $(\mathbf{P}, \mathbf{Q}, \mathbf{R})$) is winning (fair) (losing).

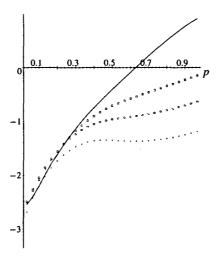


FIGURE 10. We take 3 losing games with $p_0 = p_1 \equiv p$, $q_0 = q_1 \equiv q$, and $r_0 = r_1 \equiv r$, and $p_2 = 0.8(1-p)^2/((1-p)^2+p^2)$, $q_2 = 0.8(1-q)^2/((1-q)^2+q^2)$, and $r_2 = 0.9(1-r)^2/((1-r)^2+r^2)$. We set q = 0.1 so that $q_2 = 0.8890...$, and plot $\ln(c(\mathbf{P}, \mathbf{Q}, \mathbf{R})) = d_3 + d_4$, compare (2.2), as a function of p for various values of r, solid: r = 3/4, $(r_2 = 0.9)$; box: r = 1/2, $(r_2 = 0.45)$; circles: r = 3/8, $(r_2 = 0.66176...)$; crosses r = 1/4, $(r_2 = 0.81)$. If $\ln(c(\mathbf{P}, \mathbf{Q}, \mathbf{R}))$ is positive (zero) (negative) then the game $(\mathbf{P}, \mathbf{Q}, \mathbf{R})$) is winning (fair) (losing).

shown in Figure 9, by varying one parameter three-state transitions of the composed game may be observed. This phenomenon seems to be similar to changes of the direction of the current in a multiplicative stochastic ratchet as observed in [14].

As shown in Section 3, a game constructed by composing deterministically 2-periodic fair games leads to a pseudo-paradox. In general, a game composed from an even number of 2-periodic fair games always exhibits a pseudo-paradox [13]. That is, if each simple game is 2-state periodic with an even temporal period, then the composite game is constructed only from half of the simple games, and the composite game may be winning, or losing, or fair. Moreover, when each of the simple games is fair with an odd temporal period, and the composite game is composed deterministically, then Parrondo's paradox never occurs. An example in Section 3 of a game composed from 2-state periodic and 3-temporal periodic games illustrates this statement.

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