

LETTER TO THE EDITOR

Dear Editor,

On perturbed random walks

Let (X_n, M_n) , $n \geq 1$, be an independent and identically distributed sequence of two-dimensional random vectors such that M_1 has finite mean and is long tailed (i.e. $P(M_1 > x + y)/P(M_1 > x) \rightarrow 1$ as $x \rightarrow \infty$ for y fixed), $E[X_1] = -\mu < 0$, and there exists a $\zeta > 0$ such that $E[e^{\zeta X_1}] < \infty$, i.e. the right tail of X_1 is light tailed. It was shown in [1] that

$$P\left(\sup_{n \geq 1} [M_n - \mu n] > x\right) \sim \frac{1}{\mu} \int_x^\infty P(M_1 > u) du := G_\mu(x).$$

This still holds if $-\mu n$ is replaced by $S_{n-1} = \sum_{i=1}^{n-1} X_i$. This is the main result of Ha *et al.* [3] under the additional assumption that $E[X_1^2] < \infty$; see their Theorem 1(i) in the case $\gamma = 0$. The asymptotic lower bound of this result was already covered in [4] without the assumption that $E[X_1^2] < \infty$; see the top half of page 352 of [4]. Indeed, the assumption in [4] is that $\max\{M_1, X_1\}$ is long tailed, but

$$P(M_1 > x) \leq P(\max\{M_1, X_1\} > x) \leq P(M_1 > x) + P(X_1 > x) \sim P(M_1 > x),$$

and the class of long-tailed distributions is closed under tail equivalence, so that $\max\{M_1, X_1\}$ is also long tailed.

The aim of this letter is to provide a proof of the corresponding asymptotic upper bound which is shorter and more general than the proof in [3], i.e. we will show that

$$\limsup_{x \rightarrow \infty} \frac{P(\sup_{n \geq 1} [M_n + S_{n-1}] > x)}{G_\mu(x)} \leq 1. \tag{1}$$

Define, for $\varepsilon > 0$, the event $A_\varepsilon(x) = \bigcup_n \{M_n > x + (n - 1)(\mu - \varepsilon)\}$. Also, define $\bar{S}_n = S_n + (\mu - \varepsilon)n$ and $\bar{M}_n = M_n - (\mu - \varepsilon)(n - 1)$. Then

$$P\left(\sup_{n \geq 1} [S_{n-1} + M_n] > x\right) \leq P(A_\varepsilon(x)) + \sum_{n=1}^\infty P(\bar{S}_{n-1} + \bar{M}_n > x; \bar{M}_n \leq x).$$

The first term behaves as $G_{\mu-\varepsilon}(x)$, so we focus on the second term. Since \bar{S}_1 has mean $-\varepsilon$ and has a light right tail (similarly as X_1), there exists $\theta > 0$ such that $r := E[e^{\theta \bar{S}_1}] < 1$. Thus,

$$\begin{aligned} P(\bar{S}_{n-1} + \bar{M}_n > x; \bar{M}_n \leq x) &\leq e^{-\theta x} E[e^{\theta \bar{S}_{n-1} + \bar{M}_n}; \bar{M}_n \leq x] \\ &= e^{-\theta x} r^{n-1} E[e^{\theta \bar{M}_n}; \bar{M}_n \leq x] \\ &= r^{n-1} \int_0^x e^{-\theta(x-u)} dP(\bar{M}_n \leq u) \\ &= r^{n-1} P(E_\theta + \bar{M}_n > x; \bar{M}_n \leq x) \\ &\leq r^{n-1} P(E_\theta + \bar{M}_n > x), \end{aligned}$$

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with E_θ exponential(θ) distributed. Let, in addition, N be geometric with rate r to conclude that

$$\begin{aligned} \sum_{n=1}^{\infty} \mathbb{P}(\bar{S}_{n-1} + \bar{M}_n > x; \bar{M}_n \leq x) &\leq \frac{1}{1-r} \mathbb{P}(E_\theta + N(\mu - \varepsilon) + M_1 > x) \\ &\sim \frac{1}{1-r} \mathbb{P}(M_1 > x) \\ &= o(G_\mu(x)), \end{aligned}$$

using basic properties of long-tailed distributions in the last two steps. In particular, the asymptotic equivalence follows from Breiman's [2] theorem (since the tail of e^{M_1} is slowly varying; see [2]), and the final equality follows from the elementary fact that $G_\mu(x) \geq u \mathbb{P}(M_1 > x + u)/\mu$ for every u . This implies (1) by letting ε become arbitrarily small.

This proof is shorter and more general than the proof in [3] since we do not assume that $E[X_1^2] < \infty$, we only require the right tail of X_1 be light tailed.

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Yours sincerely,

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