Effect of maternal dietary counselling during the 1st year of life on glucose profile and insulin resistance at the age of 8 years: a randomised field trial

Cintia S. Costa¹*, Paula D. B. Campagnolo², L. H. Lumey³ and Marcia R. Vitolo⁴

¹*Graduate Program in Health Sciences, Universidade Federal de Ciências da Saúde de Porto Alegre, Porto Alegre, RS, 90050170, Brazil*

²School of Health, Universidade do Vale do Rio dos Sinos, São Leopoldo, RS, 93022000, Brazil

³Department of Epidemiology, Mailman School of Public Health, Columbia University, New York, NY 10032, USA

⁴Department of Nutrition, Universidade Federal de Ciências da Saúde de Porto Alegre, Porto Alegre, RS, 90050170, Brazil

(Submitted 7 August 2016 – Final revision received 16 November 2016 – Accepted 14 December 2016 – First published online 18 January 2017)

Abstract

Education interventions that stimulate complementary feeding practices can improve the nutritional status of children and may protect against future chronic diseases. We assessed the long-term effectiveness of dietary intervention during the 1st year of life on insulin resistance levels, and investigated the relationship between insulin resistance and weight changes over time. A randomised field trial was conducted among 500 mothers who gave birth to full-term infants between October 2001 and June 2002 in a low-income area in São Leopoldo, Brazil. Mother–child pairs were randomly assigned to intervention (n 200) and control groups (n 300), and the mothers in the intervention group received dietary counselling on breast-feeding and complementary feeding of their children during the 1st year of life. Fieldworkers blinded to assignment assessed socio-demographic, dietary and anthropometric data during follow-up at ages 1, 4 and 8 years. Blood tests were performed in 305 children aged 8 years to measure fasting serum glucose and insulin concentrations and the homoeostasis model assessment index of insulin resistance (HOMA-IR). At the age of 8 years, the intervention group showed no changes in glucose and insulin concentrations or HOMA-IR values (change 0-07; 95% CI –0-09, 0-04 for boys) compared with study controls. Insulin resistance was highly correlated, however, with increases in BMI between birth and 8 years of age. Although this dietary intervention had no impact on glucose profile at age 8 years, our findings suggest that BMI changes in early childhood can serve as an effective marker of insulin resistance.

Key words: Intervention studies: Randomised controlled trials: Dietary counselling: Childhood: Insulin resistance: Homoeostasis model assessment index of insulin resistance: BMI gain

The number of overweight children under the age of 5 years in 2013 was estimated to be over forty-two million; 75% of these children are living in developing countries^(1,2). The relationship between body weight and insulin resistance in children is therefore an important research topic⁽³⁻⁵⁾. A growing body of evidence suggests that excessive weight gain in infancy and in childhood is likely to be associated with insulin resistance in adults^(6–8). The impact of early weight gain on insulin resistance in children has not been clearly identified, however. Insulin resistance in childhood is of great clinical importance as it may lead to diabetes type 2, hypertension, hepatic steatosis, endothelial dysfunction⁽⁹⁻¹¹⁾, CVD and the metabolic syndrome⁽¹²⁾ later in life. The prevalence of insulin resistance in childhood has shown a steady increase in recent years^(13,14), and this points to the need for appropriate and early preventive intervention strategies.

Systematic reviews suggest that educational interventions including breast-feeding and complementary feeding practices

may be effective in improving the nutritional status of infants and young children^(15,16), and thereby also protect against future chronic diseases⁽¹⁷⁾. We have shown in a field trial that maternal counselling for infant feeding at home can stimulate infant breast-feeding at 4, 6 and 12 months of age⁽¹⁸⁾, which could provide some long-term protection against the development of insulin resistance^(19,20). In the same field trial, maternal counselling during the 1st year of life was effective in reducing children's energy-dense food consumption at 12 months^(18,21), in improving diet quality at 4 years⁽²²⁾ and lipid profile in daughters at 8 years of $age^{(23)}$. We are not aware, however, of any studies of the potential benefits of maternal counselling on infant feeding in reducing insulin resistance in childhood, particularly in developing countries. We expect that efforts to prevent insulin resistance, especially if started early enough, could delay the progress of metabolic complications and optimise healthier outcomes⁽²⁴⁾. We therefore investigated whether the effects of our home-based infant nutrition interventions

Abbreviation: HOMA-IR, homoeostasis model assessment index of insulin resistance.

* Corresponding author: C. S. Costa, fax +55 51 3003 8798, email cintiadossantoscosta@terra.com.br

135

https://doi.org/10.1017/S0007114516004578 Published online by Cambridge University Press

were still seen after the study population had reached the age of 8 years. For this study, our focus was on selected metabolic parameters related to insulin resistance, including serum glucose, insulin and the homoeostasis model assessment index of insulin resistance (HOMA-IR) measure. To further clarify the relationship between body size and insulin resistance in children, we also examined the relationship between the metabolic parameters and the weight changes over time in children of mothers who did and did not receive dietary counselling.

Methods

The randomised field trial was conducted at the maternity wards of a hospital in a low-income population setting in the city of São Leopoldo, Brazil. Mothers of healthy, singleton, full-term (>37 weeks) and normal birth weight (\geq 2500 g) babies were invited to participate. We excluded HIV-positive mothers, infants with congenital malformations or infants who were admitted to neonatal intensive care units, and individuals with breast-feeding impediments. The present study was conducted according to the guidelines laid down in the Declaration of Helsinki, and all procedures involving human subjects were approved by the Ethics Committee of the Universidade Federal de Ciências da Saúde de Porto Alegre. All parents provided written informed consent. We followed the Consolidated Standards for Reporting Trials guidelines to report on this randomised field study⁽²⁵⁾. The study has been registered at www.clinicaltrials.gov (identifier NCT00629629).

A total of 500 mother-child pairs were recruited by fieldworkers from maternity wards, representing 89.5% of all invited mothers. To guarantee blinding of the intervention assignment, an investigator not involved in the recruitment conducted the randomisation procedure during the trial. Block randomisation was used to avoid imbalances during the randomisation process. Mothers who agreed to participate were sequentially listed on the basis of their time of delivery, grouped in blocks of five and their names separated in opaque, sealed envelopes. Two mothers from each block were randomly assigned to the intervention group, and the remaining three mothers were allocated to the control group. At the end of the randomisation, 200 children were allocated to the intervention group and 300 to the control group. We included more mother-child pairs in the control group as we expected greater losses to follow-up in controls because of a lower frequency of follow-up home visits.

For the original trial, we calculated that a sample of 363 infants would be required to detect a 65% increase in the frequency of exclusive breast-feeding up to 4 months of age in the intervention group (with 80% power and $\alpha = 5\%$), assuming a 21.5% frequency of exclusive breast-feeding in the control group. For the current follow-up at age 8 years, 128 children would be required to detect a HOMA-IR change of 0.5 sp units, with 80% power and $\alpha = 5\%$.

Intervention

The intervention consisted of dietary advice about breast-feeding and complementary feeding based on the 'Ten steps for the healthy feeding for Brazilian children from birth to two years of age'(26). It was carried out between October 2001 and June 2002 by home visits within 10d of the child's birth, on a monthly basis up to 6 months of age, and at 8, 10 and 12 months of age. The main purpose of the programme was to promote exclusive breast-feeding for 6 months followed by healthy complementary foods. During each home visit, mothers received dietary advice in accordance with the baby's age. Mothers were advised against the addition of sugars (cane sugar, honey) to fruit, porridge, juices, milk or other liquids. They were encouraged to avoid fried food, soft drinks, sweets and salty snacks and to use salt in moderation. Advice on hygiene practices in food preparation and handling was provided. A simple coloured leaflet with food pictures comprising a healthy meal was used to guide the dietary advice and was given to the mother as a reminder. The writing material was simplified to take into consideration the mothers' level of education. During each visit of about 40-60 min, the fieldworkers clarified and reinforced recommendations while respecting the mother's level of cognition as well as cultural and economic background. The dietary intervention summary and the main counselling strategies applied during each home visit have been described elsewhere in detail⁽²³⁾. The counselling was carried out by paired undergraduate students in nutrition science. The fieldworkers who carried out the dietary advice received 8h of theoretical training. During the intervention programme, quality control was ensured by weekly scheduled meetings with all fieldworkers and the coordinator of the programme to discuss all dietary advice provided to mothers. Mothers were encouraged to report any adverse events that occurred with children during the intervention.

Control group

Mothers in the control group received the recommended standard care. They were interviewed twice during the 1st year after childbirth (at 6 and 12 months of age) for data collection only. All mothers were encouraged to maintain normal paediatric visits for their babies during the study period. Nutritional diagnoses were provided to mothers, and they were advised to talk to the health professionals about the nutritional diagnosis that we provided to them. After the 1st year, children in intervention and control groups were followed-up during childhood, and new sets of data were collected at 4 and 8 years of age.

Data collection

Data for identification and data required for locating the families in the community were collected at the time of recruitment. Trained fieldworkers, not involved in the intervention and who were unaware of group allocation, conducted face-to-face structured interviews during home visits with the mothers at 1, 4 and 8 years following the infants' birth. Every month, 10% of the questionnaires were selected randomly and followed-up by telephone calls to the mothers to verify the authenticity of the collected data. Children's sex, skin colour, birth weight and mode of delivery were obtained from hospital records. Prepregnancy weight and maternal weight at the end of pregnancy

136

were self-reported, and mothers' height was measured during home visits by fieldworkers when the children were 6 months old. Pre-pregnancy BMI was calculated as pre-pregnancy weight divided by the square of height (kg/m²). Gestational weight gain was calculated by subtracting pre-pregnancy weight from weight at the end of pregnancy. Household income and duration of exclusive breast-feeding data were collected during home visits. When the children were 1-year old, the mothers were asked whether their infants had received regular healthcare services in the 1st year after birth. At 8 years of age, for diet pattern analyses, two multiple-pass 24-h dietary recalls were collected for each child on two randomly selected and non-consecutive days upon home visits to the families. The mean nutritional composition of the two 24-h dietary recalls for each child was classified according to the Healthy Eating Index (HEI)⁽²⁷⁾, which is an instrument that attributes scores to the diet quality of individuals; the details have been described elsewhere^(28,29). In order to assess a sedentary lifestyle marker among 8-year-old children, mothers were asked to report the total (hours and minutes) night-time sleep duration and the total (hours and minutes) screen-time duration (including television, computer and video-game) on the preceding day of the interview.

Anthropometric measurements

At 1 year of age, all children were weighed without clothing on a portable digital scale (Techline), and their length was measured by using an infant stadiometer (Serwital Inc.). At ages 4 and 8 years, children were weighed to the nearest 0·1 kg in light clothing without shoes on a digital scale, and standing height was measured to the nearest 0·1 cm using a stadiometer (SECA). All measures were converted into *z*-scores of BMI-for-age on the basis of World Health Organization Growth Standards^(30,31). Changes in growth measurements from birth to 8 years of age were analysed as BMI-for-age *z*-score variation, considering three periods: from birth to 1 year, from 1 to 4 years and from 4 to 8 years of age.

Glucose profile

At 8 years of age, venous blood samples were obtained from the right arm after an overnight fast to measure serum glucose and insulin concentrations and to calculate the HOMA-IR index. Analyses were performed at the laboratory of Cardiology Institute of Rio Grande do Sul by technicians who were unaware of study assignments. Glucose and insulin were estimated using an automatic analyser (Cobas Integra[®]; Roche). HOMA-IR was calculated as (insulin (μ U/ml) × glucose (mmol/l))/22·5⁽³²⁾.

Statistical analyses

Analyses were performed by intention-to-treat and by sex. Non-normally distributed variables were log transformed for all statistical procedures. Untransformed values are presented in all tables for ease of clinical interpretation. Student's *t* test was used to evaluate the effect of the intervention on independent continuous variables. Univariate and multivariate linear regressions

were performed to examine the relation of intervention and selected anthropometric covariates of special interest on glucose and insulin concentration and HOMA-IR values at age 8 years, including pre-pregnancy BMI, gestational weight gain, child's birth weight and BMI z-score from birth to 1, 4 and 8 years of age. In univariate models, the reported β -coefficients represent separate models for each listed variable, and in multivariate models β -coefficients are adjusted for all variables in the model. We also examined the effects of adjustment for baseline social, demographic and breast-feeding variables child's sex (male/female) and skin colour (white/not-white), maternal schooling, total household income, mode of delivery (normal/caesarean), exclusive breast-feeding (≥ 4 months) and breast-feeding - on our estimates of treatment effects. Finally, we further analysed the effect of adjustment for total HEI score, total sleep and screen-time duration at 8 years of age on our estimates of treatment effects. Collinearity was checked in all models. All statistical analyses were performed using SPSS 16.0 (SPSS IBM Inc.), and statistical significance was set at P < 0.05(two-sided).

Results

Among the 500 children initially recruited, 396 underwent anthropometric assessments at age 1 year, 345 at age 4 years and 309 at age 8 years (Fig. 1). A total of 305 children underwent glucose assessment and 303 children underwent insulin and HOMA-IR assessment at 8 years of age. No adverse events were reported during the intervention. The proportion of overweight children (BMI>1 sD) was 36.1% (n 143) at age 1 year, 20.6% (*n* 71) at age 4 years and 27.5% (*n* 85) at age 8 years. There were no differences in overweight prevalence proportions between intervention and control groups for the three periods (1, 4 and 8 years), and this result persisted after analyses by sex. The median duration of exclusive breast-feeding was 3.5 months (95% CI 0.5, 6.5) in the intervention group and 1.5 months (95% CI 0.5, 6.5) in the control group; the median duration of breast-feeding was 12.5 months (95% CI 0.5, 12.5) in the intervention group and 10.5 months (95% CI 0.5, 12.5) in the control group.

No differences were found between children who were lost to follow-up and those who remained at 8 years of age in terms of pre-pregnancy BMI (P=0.48), gestational weight gain (P=0.89), mode of delivery (P=0.88), weight at birth (P=0.55), maternal age at child's birth (P=0.22) and maternal level of education (P=0.66). There were no differences between intervention and control groups on selected baseline characteristics (Table 1).

There were no differences between intervention and control groups with respect to glucose and insulin concentrations and HOMA-IR indices at 8 years of age, for both sexs (Table 2). In addition, there were no significant differences between intervention and control groups comparing BMI *z*-score changes from birth to 1 year, from age 1 to 4 years and from age 4 to 8 years, in girls and boys.

The linear regression analyses evaluating the associations between anthropometric covariates and glucose, insulin and MS British Journal of Nutrition

https://doi.org/10.1017/S0007114516004578 Published online by Cambridge University Press





HOMA-IR values at 8 years of age are shown in Table 3. In multivariate analysis, birth weight and the increase in BMI *z*-scores between 1 and 4 years and between 4 and 8 years all contributed significantly to glucose and insulin concentrations at age 8 years and to rises in HOMA-IR. The increase in BMI during the 1st year of life was positively associated with rises in insulin and HOMA-IR. These findings persisted after adjustment for child sex and skin colour, total family income, maternal schooling, mode of delivery and exclusive or partial breastfeeding. Additional analyses were performed after further adjustment for total sleep and screen-time duration and total HEI score at 8 years of age and all results persisted.

Discussion

To the best of our knowledge, this is the first randomised trial to examine the potential effects on children of maternal dietary counselling during the 1st year of life on insulin resistance at age 8. Contrary to our expectation, we did not see any changes in insulin resistance at this age. Table 1. Characteristics of children and their households at baseline according to the group, at age 1 year* (Numbers and percentages; mean values and standard deviations)

		Intervention		Control		
Characteristics	n*	п	%	n	%	<i>P</i> †
Child						
Boys	397	93	57.1	131	56.0	0.83
Skin colour (white)	331	58	41.1	85	44.7	0.51
Birth weight (g)	390					0.66
Mean		3375-34		3354-30		
SD		464.51		466.70		
Birth length (cm)	390					0.62
Mean		4	8.74	4	8.84	
SD			1.92		2.12	
Households at baseline						
Maternal pre-gestational BMI (kg/m ²)	367					0.10
Mean		2	4.72	2	3.90	
SD		5.32		4.15		
Gestational age (weeks)	375					0.76
Mean		39.34		39.38		
SD		1.31		1.19		
Delivery by caesarean section	356	51	36.7	94	43.3	0.21
Maternal age at child's birth <20 years	397	29	17.8	46	19.7	0.64
Mother's education <8 years	396	99	61.1	131	56.0	0.31
Father's education <8 years	369	93	61.6	120	55.0	0.21
Mother's employment	391	58	35.8	76	33.2	0.59
Father's employment	361	130	88.4	195	91·1	0.40
Annual household income ≤U\$300	395	122	74.8	165	71.1	0.18

* *n* Indicates the number of responses recorded for each characteristic.

+ Student's t test or χ^2 was used (depending on categorical or continuous variables).

 Table 2. Intervention effect on glucose, insulin and homoeostasis model assessment index of insulin resistance (HOMA-IR)* at 8 years of age and on BMI z-score variation from birth to 8 years of age, according to sex

 (Mean values and standard deviations; differences and 95% confidence intervals)

	Intervention				Control				
	n	Mean	SD	n	Mean	SD	Difference	95 % CI	Ρ
Girls									
Glucose (mmol/l)†	52	4.22	0.42	80	4.24	0.38	0.02	- 0·12, 0·16	0.80
Insulin (µU/ml)†‡	52	5.57	3.12	79	7.32	7.02	0.07	-0.06, 0.21	0.28
HOMA-IR†‡	52	1.06	0.62	79	1.41	1.40	0.07	-0.06, 0.21	0.29
∆BMI z-score birth to 1 year†	70	0.17	1.59	99	0.28	1.18	0.11	-0.31, 0.53	0.61
∆BMI z-score from 1 to 4 years†	60	-0.44	0.82	86	-0.45	1.00	-0.01	-0.32, 0.30	0.94
∆BMI z-score 4–8 years†	49	0.04	0.87	79	0.19	0.82	0.15	-0.14, 0.45	0.32
Boys									
Glucose (mmol/l)†	74	4.46	0.42	99	4.46	0.42	-0.01	- 0·13, 0·12	0.97
Insulin (µU/ml)†‡	73	5.74	4.32	99	4.53	3.14	-0.07	-0.19, 0.03	0.17
HOMA-IR†‡	73	1.15	0.87	99	0.92	0.62	-0.07	-0.19, 0.04	0.17
∆BMI z-score birth to 1 year†	93	0.10	1.22	127	0.03	1.25	-0.07	- 0·40, 0·27	0.69
∆BMI z-score from 1 to 4 years†	82	-0.21	1.32	110	-0.28	1.05	-0.07	- 0·41, 0·27	0.67
ΔBMI z-score 4-8 years†	72	0.24	1.11	93	-0.06	1.18	- 0.29	-0.65, 0.06	0.10

* HOMA-IR: (insulin (μU/ml) × glucose (mmol/l)/22·5).

† Student's t test was used.

‡ Non-normally distributed variables were log transformed.

We considered two possibilities that may explain the lack of effectiveness of the trial to affect metabolic outcomes in childhood. First, the change in dietary practices observed in this population^(18,21,22) might not be large enough to have any long-term impact on glucose metabolism at age 8 years. This supports the need to continue the dietary counselling even after the infancy period, to achieve adequate public policies. Second, any early changes in dietary practices may have been overridden by dietary and weight gain changes between 1 and 8 years of age, after the study intervention had ceased at age 1 year. The focus of the 'Ten steps for the healthy feeding for Brazilian children from birth to two years of age' intervention plan is on dietary counselling, and monitoring excessive weight gain is not emphasised by the plan. As insulin resistance is closely related to body weight^(3,4,33), the lack of effectiveness of the trial to affect metabolic outcomes in older children could be related to weight gains between 1 and 8 years of age in all study participants, irrespective of dietary counselling in the

https://doi.org/10.1017/S0007114516004578 Published online by Cambridge University Press

Table 3. Linear regressions analysis of glucose, insulin and homoeostatic model assessment of insulin resistance (HOMA-IR)* at 8 years of age and independent variables (*n* 305) (β-Coefficients and 95% confidence intervals)

Univariate model† Multivariate model± Р Р β 95 % CI β 95 % CI Glucose (mmol/l) 0.00 -0.10.0.10 0.98 -0.02 0.12.0.08 0.71 Intervention Pre-pregnancy BMI (kg/m²) 0.17-0.03, 0.36 0.09-0.03 -0.24, 0.18 0.77 Gestational weight gain (kg) 0.03 -0.11, 0.18 0.63 -0.02 -0.18, 0.13 0.77 Birth weight (kg) 1.11 -0.78.3.00 0.25 2.79 0.44.5.14 0.02 ΔBMI z-score birth to 1 year 0.25 -0.94, 0.43 0.46 0.64 -0.24, 1.52 0.15 ABMI z-score 1-4 years 1.17 0.36 1.99 0.01 1.46 0.51 2.41 0.01 ∆BMI z-score 4-8 years 1.61 0.78, 2.43 <0.001 1.79 0.89, 2.70 <0.001 Insulin (µU/ml) (log) -0.03 Intervention 0.01 -0.08, 0.10 0.81 -0.11, 0.070.56 Pre-pregnancy BMI (kg/m²) 0.01 0.00, 0.02 0.04 0.00 -0.01, 0.01 0.98 Gestational weight gain (kg) -0.01.0.01 0.72 - 0.00 -0.01.0.00 0.47 - 0.00 Birth weight (kg) 0.00 -0.09, 0.10 0.98 0.12 0.01, 0.24 0.03 ΔBMI z-score birth to 1 year -0.02, 0.04 0.04, 0.12 <0.001 0.01 0.48 0.08 ∆BMI z-score 1-4 years 0.08 0.04 0.12 < 0.001 0.12 0.08 0.17 < 0.001 ∆BMI z-score 4-8 years 0.09 0.05, 0.13 <0.001 0.11 0.07, 0.15 <0.001 HOMA-IR (log) 0.01 -0.08, 0.10 0.82 -0.03-0.12, 0.06 0.54 Intervention 0.00, 0.02 Pre-pregnancy BMI (kg/m²) 0.03 -0.01, 0.01 0.01 0.00 0.99 Gestational weight gain (kg) -0.01.0.01 - 0.00 0.77 0.00 -0.01 0.00 0.45 Birth weight (kg) 0.01 -0.09, 0.11 0.88 0.14 0.02, 0.30 0.02 ΔBMI z-score birth to 1 year 0.01 -0.03, 0.04 0.55 0.08 0.04, 0.18 <0.001 ∆BMI z-score 1-4 years 0.09 0.05, 0.13 <0.001 0.13 0.04, 0.14 <0.001 ∆BMI z-score 4-8 years 0.10 0.06. 0.14 <0.001 0.12 0.05. 0.14 <0.001

* HOMA-IR: (insulin (μU/ml)×glucose (mmol/l)/22·5).

† Univariable model: β-coefficients for a separate model for each listed variable.

‡ Multivariate model adjusted for all variables together.

1st year of life. Although the effectiveness of breast-feeding and complementary feeding interventions to improve child nutrition, growth and development is well documented in many settings^(34–37), less attention has been paid in clinical trials to the potential impact of weight gain in infancy and childhood as a contributor to the development of adverse metabolic outcomes, including insulin resistance in children.

In the absence of an intervention effect, we therefore also analysed the impact of selected anthropometric variables on the glucose profile and insulin resistance at ages 1, 4 and 8 years. Our findings show that increased birth weight and higher weight gain velocities from birth to 8 years of age all had a negative impact on the glucose profile and insulin resistance among school-aged children, irrespective of the dietary counselling in the 1st year of life. These findings support other observations that higher weight gain velocities at either 3-5 years⁽³⁸⁾ or 4-9 years of age⁽³⁹⁾ can lead to relatively larger increases in fat mass and insulin resistance, either through overproduction of NEFA or through increased synthesis and release of pro-inflammatory cytokines⁽⁴⁰⁾. Recently, a positive relationship was observed between BMI z-score increases and adverse insulin and HOMA-IR outcomes in Brazilian children followed-up from 4 to 10 years, although no impact was found on glucose concentrations⁽⁴¹⁾. Weight gain from birth to 2 years of age is already related to insulin resistance among young adults in several cohort studies conducted in developing countries⁽⁴²⁾. There is also evidence that increases in fasting insulin between 3 and 6 years of age can be associated with a greater risk of type 2 diabetes,

irrespective of adult BMI and parental history⁽⁴³⁾. Although major metabolic diseases related to early insulin resistance may not become symptomatic until adulthood, our findings suggest that the detection and prevention of risk factors for insulin resistance should begin in childhood, when changes in lifestyle, including those related to weight gain, can still reduce the risk and severity of metabolic disease later in life.

This study has some limitations. Although we experienced losses to follow-up, we found no differences between the baseline characteristics of children who remained in the study and those who were lost to follow-up. Second, the mothers may have been aware of the intervention group to which they were assigned. This may have affected their responses to the study because of social desirability bias, as it is not possible to blind patients in studies that evaluate dietary advices. To minimise such bias, fieldworkers not involved in the intervention carried out the assessments. A limitation with respect to generalisability of the study findings is the choice of the study population, including only children between birth and 8 years of age from a low-income population. In response, we note that this target population is of specific interest for the introduction and qevaluation of large-scale dietary interventions to improve population health in regional or national health programmes. There could be self-report bias of pre-pregnancy weight: however, studies conducted in high-⁽⁴⁴⁾ and low-income countries⁽⁴⁵⁾, including a study in Brazil⁽⁴⁶⁾, have demonstrated that the BMI obtained using self-reported pre-pregnancy weight strongly correlated with those obtained using

anthropometrics measured. Considering the outcome of this study (insulin resistance), another limitation is the fact that gestational diabetes data were not obtained and children born from those mothers could present metabolic alterations. However, this trial was conducted with healthy newborn babies, and the analyses were adjusted for pre-pregnancy weight, gestational weight gain and birth weight - variables highly related to gestational diabetes mellitus. As a further limitation, we were not able to measure insulin resistance before age 8 years, and other studies will be needed to fill this gap. This study clearly shows, however, that understanding the metabolic impact of dietary interventions and growth trends in infancy and childhood is extremely important to formulate public health strategies aimed at preventing chronic diseases and that the long-term follow-up even of study populations recruited at birth can be extremely important.

In summary, our study shows no impact of dietary counselling in the 1st year of life on metabolic profiles and insulin resistance at the age of 8 years. Our results do point, however, to the crucial relationship between infant and childhood weight gain and insulin resistance at age 8 years in school children from a low-income community in Brazil. We found that weight gains between birth and ages 1, 1–4 and 4–8 years all contributed significantly to adverse changes in insulin resistance at the age of 8 years. Our findings suggest that preventing excessive weight gain since early life is a relevant key to prevent insulin resistance in childhood.

Acknowledgements

The authors thank the families who participated in the study. The present study was supported by the Brazil CNPq (National

Council for Scientific and Technological Development).

C. S. C.: formulated the research question, analysed and interpreted the data, performed statistical analysis and wrote the manuscript; P. D. B. C.: conducted the study and critically reviewed the manuscript; L. H. L.: interpreted the data and statistical analysis and critically reviewed the manuscript; M. R. V.: designed and conducted the study, formulated the research question, interpreted the data and the statistical analysis, and critically reviewed the article. All authors: read and approved the final manuscript.

The authors declare that there are no conflicts of interest.

References

- World Health Organization (2016) Report of the Commission on Ending Childhood Obesity. http://apps.who.int/iris/bitstream/10665/204176/1/9789241510066_eng.pdf (accessed June 2016).
- 2. Ng M, Fleming T, Robinson M, *et al.* (2014) Global, regional, and national prevalence of overweight and obesity in children and adults during 1980-2013: a systematic analysis for the Global Burden of Disease Study 2013. *Lancet* **30**, 766–781.
- Henderson M, Benedetti A, Barnet TA, *et al.* (2016) Influence of adiposity, physical activity, fitness, and screen time on insulin dynamics over 2 years in children. *JAMA Pediatrics* 8, e1–e9.
- Santiago-Torres M, Cui Y & Adams AK (2016) Familial and individual predictors of obesity and insulin resistance in urban Hispanic children. *Pediatr Obes* 11, 54–60.

- Masquio DCL, Piano A, Campos RMS, *et al.* (2015) The role of multicomponent therapy in the metabolic syndrome, inflammation and cardiovascular risk in obese adolescentes. *Br J Nutr* **113**, 1920–1930.
- Barghava SK, Sachdev HS, Fall CHD, *et al.* (2004) Relation of serial changes in childhood body-mass index to impaired glucose tolerance in young adulthood. *N Engl J Med* 350, 865–875.
- Fabricius-Bjerre S, Jensen RB, Faerch K, *et al.* (2011) Impact of birth weight and early infant weight gain on insulin resistance and associated cardiovascular risk factors in adolescence. *PLoS ONE* 6, 1–8.
- Slining MM, Kuzawa CW, Mayer-Davis EJ, *et al.* (2011) Evaluating the indirect effect of infant weight velocity on insulin resistance in young adulthood: a birth cohort study from the Philippines. *Am J Epidemiol* **173**, 640–648.
- 9. Ho M, Garnett SP & Baur LA (2014) Childhood obesity and insulin resistance: how should it be managed? *Curr Treat Options Cardiovasc Med* **16**, 351–367.
- Nobili V, Alkhouri N, Alisi A, *et al.* (2015) Nonalcoholic fatty liver disease: a challenge for pediatricians. *JAMA Pediatr* 169, 170–176.
- Sartorio A, Del Col A, Agosti F, *et al.* (2007) Predictors of non-alcoholic fatty liver disease in obese children. *Eur J Clin Nutr* 61, 877–883.
- Ten S & Maclaren N (2010) Insulin resistance syndrome in children. J Clin Endocrinol Metab 89, 2526–2539.
- Chiarelli F & Marcovecchio ML (2008) Insulin resistance and obesity in childhood. *Eur J Endocrinol* 159, s67–s74.
- Van Der Aa MP, Farsini SF, Knibbe CAJ, et al. (2015) Population-based studies on the epidemiology of insulin resistance in children. J Diabetes Res 2015, 1–9.
- Shi L & Zhang J (2011) Recent evidence of the effectiveness of educational interventions for improving complementary feeding practices in developing countries. *J Trop Pediatr* 57, 91–98.
- Imdad A, Yakoob MY & Bhutta ZA (2011) Impact of maternal education about complementary feeding and provision of complementary foods on child growth in developing countries. *BMC Public Health* **11**, Suppl. 3, 525–539.
- 17. Lanigan J & Singhal A (2009) Early nutrition and long-term health: a practical approach. *Proc Nutr Soc* **68**, 422–429.
- Vitolo MR, Bortolini GA, Feldens CA, *et al.* (2005) Impacts of the 10 steps to healthy feeding in infants: a randomized field trial. *Cad Saude Publica* 21, 1448–1457 (in Portuguese).
- Veena SR, Krishnaveni GV, Wills JC, et al. (2011) Glucose tolerance and insulin resistance in indian children: relationship to infant feeding pattern. *Diabetologia* 54, 2533–2537.
- Owen C, Martin RM & Whincup PH (2006) Does breastfeeding influence risk of type 2 diabetes in later life? A quantitative analysis of published evidence. *Am J Clin Nutr* 84, 1043–1054.
- Vitolo MRV, Bortolini GA, Campagnolo PDB, *et al.* (2012) Maternal dietary counseling reduces consumption of energydense foods among infants: a randomized controlled trial. *J Nutr Educ Behav* 44, 140–147.
- 22. Vitolo MRV, Rauber F, Campagnolo PDB, *et al.* (2010) Maternal dietary counseling in the first year of life is associated with a higher healthy eating index in childhood. *J Nutr* **140**, 2002–2207.
- Louzada MLC, Campagnolo PDB, Rauber F, *et al.* (2012) Long-term effectiveness of maternal dietary counseling in a low-income population: a randomized field trial. *Pediatrics* 129, 1–8.
- 24. Zheng J, Xiao X, Zhang Q, *et al.* (2014) DNA methylation: the pivotal interaction between early-life nutrition and glucose metabolism in later life. *Br J Nutr* **112**, 1850–1857.

140

- Schulz KF, Altman DG & Moher D (2010) CONSORT 2010 Statement: updated guidelines for reporting parallel group randomized trials. *BMC Med* 24, 8–18.
- 26. Ministry of Health (2002). Dez passos para uma alimentação saudável: guia alimentar para menores de dois anos (Ten Steps for Healthy Feeding: A Guide for Children Less than Two Years of Age). Brasília: Ministry of Health (in Portuguese).
- 27. Kennedy ET, Ohls J, Carlson S, *et al.* (1995) The Healthy Eating Index: design and applications. *J Am Diet Assoc* **95**, 1103–1138.
- 28. Rauber F, Louzada MLC, Feldens CA, *et al.* (2012) Maternal and family characteristics associated with the Healthy Eating Index among low socioeconomic status Brazilian children. *J Hum Nutr Diet* **26**, 369–379.
- Rauber F, Louzada MLC & Vitolo MR (2014) Healthy eating index measures diet quality of Brazilian children of low socioeconomic status. J Am Coll Nutr 33, 26–31.
- World Health Organization Multicentre Growth Reference Study Group (2006) WHO child growth standards based on length/ height, weight and age. *Acta Paediatr*, Suppl. 450, 76–85.
- de Onis M, Onyango AW, Borghi E, *et al.* (2007) Development of a WHO growth reference for school-aged children and adolescents. *Bull World Health Organ* 85, 660–667.
- 32. Matthews DR, HoskeR JP, Rudenski AS, *et al.* (1985) Homeostasis model assessment: insulin resistance and beta-cell function from fasting plasma glucose and insulin concentrations in man. *Diabetologia* 28, 412–419.
- 33. Nakasone Y, Miyakoshi T, Sato Y, *et al.* (2016) Impact of weight gain on the evolution and regression of prediabetes: a quantitative analysis. *Eur J Clin Nutr* **70**, 757–765.
- Dewey KG & Adu-Afarwuah S (2008) Systematic review of the efficacy and effectiveness of complementary feeding interventions in developing countries. *Matern Child Nutr* 4, 24–85.
- Kramer MS & Kakuma R (2012) Optimal duration of exclusive breastfeeding. *The Cochrane Database of Systematic Reviews*, issue 8, CD003517.
- 36. Imdad A, Yakoob MY & Brutta ZA (2011) Impact of maternal education about complementary feeding and provision of complementary foods on child growth in developing countries. *BMC Public Health* **11**, 2–14.

- 37. Lassi ZS, Das JK, Zahid G, *et al.* (2013) Impact of education and provision of complementary feeding on growth and morbidity in children less than 2 years of age in developing countries: a systematic review. *BMC Public Health* **13**, 1–10.
- 38. Botton J, Heude B, Maccario J, *et al.* (2008) Postnatal weight and height growth velocities at different ages between birth and 5 y and body composition in adolescent boys and girls. *Am J Clin Nutr* 87, 1760–1768.
- Wells JC, Hallal PC, Wright A, *et al.* (2005) Fetal, infant and childhood growth: relationships with body composition in Brazilian boys aged 9 years. *Int J Obes* 29, 1192–1198.
- Koyama S, Ichikawa G, Kojima M, *et al.* (2014) Adiposity rebound and the development of metabolic syndrome. *Pediatrics* 133, e114–e119.
- Lourenço BH, Gimeno SGA & Cardoso MA (2014) BMI gain and insulin resistance among school-age children: a population-based longitudinal study in Brazilian Amazon. *Br J Nutr* 12, 1905–1910.
- Norris SA, Osmond C, Gigante D, *et al.* (2012) Size at birth, weight gain in infancy and childhood, and adult diabetes risk in five low- or middle-income country birth cohorts. *Diabetes Care* 35, 72–79.
- Sabin MA, Magnussen CG, Juonala M, *et al.* (2015) Insulin and BMI as predictors of adult type 2 diabetes mellitus. *Pediatrics* 135, e144–e151.
- 44. Han E, Abrams B, Sridhar S, *et al.* (2016) Validity of self-reported pre-pregnancy weight and body mass index classification in an integrated health care delivery system. *Paediatr Perinatal Epidemiol* **30**, 314–319.
- 45. Natamba BK, Sanchez SE, Gelaye B, *et al.* (2016) Concordance between self-reported pre-pregnancy body mass index (BMI) and BMI measured at the first prenatal study contact. *BMC Pregnancy Childbirth* **16**, 187–194.
- 46. Oliveira AF, Gadelha AMJ, Leal M, *et al.* (2004) Study of validity in self-reported weight and height among pregnant women treated at municipal maternity hospitals in Rio de Janeiro, Brazil. *Cad Saúde Pública* **20**, 92–100.