Inflammatory potential of diet and risk for hepatocellular cancer in a case-control study from Italy

Nitin Shivappa^{1,2}*, James R. Hébert^{1,2}, Jerry Polesel³, Antonella Zucchetto³, Anna Crispo⁴, Maurizio Montella⁴, Silvia Franceschi⁵, Marta Rossi⁶, Carlo La Vecchia⁶ and Diego Serraino³

¹Cancer Prevention and Control Program, University of South Carolina, Columbia, SC 29208, USA ²Department of Epidemiology and Biostatistics, Arnold School of Public Health, University of South Carolina, Columbia, SC 29208, USA

³Epidemiology and Biostatistics Unit, CRO Aviano, National Cancer Institute, 33081 Aviano, Italy

⁴Department of Epidemiology, 'Fondazione G. Pascale', Istituto Nazionale Tumori, 80133 Naples, Italy

 5 Infections and Cancer Epidemiology Group, International Agency for Research on Cancer, Lyon 69372, France

⁶Department of Clinical Sciences and Community Health, Università degli Studi di Milano, via G. Venezian 1, 20133 Milan, Italy

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Abstract

Inflammation and diet have been suggested to be important risk factors for hepatocellular cancer (HCC). This Italian multicentre hospitalbased case–control study conducted between 1999 and 2002 and including 185 cases with incident, histologically confirmed HCC, and 404 controls hospitalised for acute non-neoplastic diseases provided an opportunity to investigate the association between HCC and the dietary inflammatory index (DII). The DII was computed on the basis of dietary intake assessed 2 years before the date of interview by a validated sixty-three-item FFQ. Logistic regression models were used to estimate OR adjusted for age, sex, study centre, education, BMI, smoking, physical activity, serum markers of hepatitis B and C infection and energy intake. Energy adjustment for DII was performed using the residual method. Participants in the highest tertile of DII scores (i.e. with a more pro-inflammatory diet) had a higher risk for HCC (OR_{tertile 3 v. 1} 2·43; 95 % CI 1·27, 4·68; P_{trend} = 0·03). When stratified by the presence or absence of hepatitis B/C infection and sex, DII was strongly associated with HCC in hepatitis B- and C-negative participants (OR_{tertile 3 v. 1} 4·18; 95 % CI 1·53, 11·39; P_{trend} = 0·02) and among males (OR_{tertile 3 v. 1} 3·60; 95 % CI 1·65, 7·87; P_{trend} = 0·001). These results indicate that a pro-inflammatory diet is associated with increased risk for HCC, in those without a history of hepatitis B/C infection and among males.

Key words: Diets: Inflammation: Hepatocellular cancer: Italy: Dietary surveys and nutritional epidemiology

Hepatocellular cancer (HCC) is the most frequent primary liver tumour and the second most lethal among all human neoplasms⁽¹⁾. Approximately 80 % of HCC occur in developing countries of Asia and Africa. However, incidence of HCC and mortality from this malignancy have been increasing in the USA and parts of Europe (including Italy, until the mid-1990s)^(2–4). Inflammation is the body's response to any kind of tissue injury or insult in the presence of inflammatory stimulants such as cytokines^(5,6). Chronic inflammation – which is characterised by the continuous presence of inflammatory cytokines in circulation and in the tissues – is known to play a key role in the development of various cancers, including HCC⁽⁷⁾ where chronically inflamed liver parenchyma represents a precancerous milieu in which 70–90 % of the HCC arise⁽⁸⁾. The key causes of HCC are hepatitis B and C infection, which lead to chronic liver inflammation. Other important risk factors for HCC include smoking, alcohol, obesity and diabetes^(9,10).

Besides hepatitis B and C, there is growing evidence that specific dietary components influence both inflammation⁽¹¹⁻¹⁴⁾ and HCC⁽¹⁵⁻¹⁹⁾. Research on the role of diet in inflammation suggests that diet represents a complex set of exposures that often interact and whose cumulative effect modifies both inflammatory responses and health outcomes. A literature-derived, population-based dietary inflammatory index (DII) was developed to assess the inflammatory potential of an individual's diet⁽²⁰⁾. A pro-inflammatory diet is rich in consumption of food items rich in SFA, carbohydrate and protein and low in the consumption of PUFA, flavonoids and other dietary components. The DII has been validated with various inflammatory markers, including C-reactive protein (CRP)⁽²¹⁾

Abbreviations: DII, dietary inflammatory index; HCC, hepatocellular cancer.

^{*} Corresponding author: Dr N. Shivappa, email shivappa@mailbox.sc.edu

and IL-6^(22,23). Additionally, the DII has been shown to be associated with glucose intolerance and dyslipidaemia components of the metabolic syndrome^(24,25), anthropometric measurements in Spain⁽²⁶⁾, asthma in Australia⁽²³⁾, bone mineral density among postmenopausal women in Iran⁽²⁷⁾, colorectal cancer in a case–control study in Spain⁽²⁸⁾ and in two cohort studies of women in the USA^(29,30) and pancreatic and prostate cancers in Italy^(31,32).

Our hypothesis is that a higher DII score (indicating a pro-inflammatory diet) may increase the risk for incident HCC. In the current study, we thus examined the association between DII and HCC using a multicentre, hospital-based case–control study conducted in Italy⁽³³⁾. This provided original information on a southern European population in which hepatitis B and C are common and dietary and lifestyle habits and awareness of diet-related health issues are different from those in North America and northern Europe.

Methods

Recruitment and questionnaire

A case-control study of HCC was conducted between January 1999 and July 2002 in the province of Pordenone in Northeast Italy and the city of Naples in South Italy⁽³³⁾. Cases were 258 patients under the age of 85 years with incident HCC, who had not yet received any cancer treatment at study entry. They were admitted to Centro di Riferimento Oncologico (CRO), National Cancer Institute, Aviano, to 'Santa Maria degli Angeli' General Hospital, Pordenone and to IRCCS 'Pascale' National Cancer Institute and four General Hospitals in Naples. Overall, twenty-nine cases did not provide a blood sample and forty-four did not provide data on dietary habits, thus leaving 185 eligible cases (median age 66, range 43-84 years) with available questionnaires and blood samples for the present analysis. Histologic or cytologic confirmation was available for 78.2% of HCC cases, whereas for the remaining cases the diagnosis was based on ultrasound, tomography and elevated α -fetoprotein levels.

Controls were patients <85 years of age admitted for a wide spectrum of acute conditions to the same hospitals where HCC cases had been interviewed. Patients whose hospital admissions were due to diseases related to tobacco smoking or alcohol abuse were specifically excluded, as were those hospitalised for chronic diseases that might have led to substantial dietary modifications. However, comorbidity for such diseases was not an exclusion criterion. Blood samples were available for 431 of 462 controls; of these, 404 provided comprehensive questionnaire information on dietary habits and were included in the present analyses (median age 65, range 40-82 years). A total of 27% was admitted for trauma, 24% for non-traumatic orthopaedic diseases, 25% for acute surgical conditions, 13% for eye diseases and 11 % for other miscellaneous illnesses. Overall, 1 % of cases and controls contacted refused to participate. All study participants signed an informed consent form, according to the recommendations of the Ethical Committee of CRO, the National Cancer Institute at Aviano. Cases and controls in each study centre were interviewed in the hospital by uniformly

trained personnel using a standardised, structured questionnaire designed to collect information on socio-demographic characteristics, lifestyle habits – such as tobacco smoking and alcohol drinking – and personal medical history, including history of cirrhosis and diabetes mellitus.

An interviewer-administered FFQ shown to have satisfactory reproducibility and validity^(34,35) was used to assess participants' habitual diet, including total energy. Average weekly frequency of consumption of sixty-three foods or food groups, as well as complex recipes, during the 2 years before cancer diagnosis or hospital admission (for controls) was assessed. To compute energy and nutrient intakes, an Italian food composition database, as well as information from additional sources, was used⁽³⁶⁾.

BMI was calculated as weight (kg) divided by height squared (m²) and was categorised into normal weight (BMI < 25.0 kg/m^2), overweight ($25.0 \le BMI < 30.0 \text{ kg/m}^2$) and obese (BMI ≥ 30.0 kg/m^2).

Dietary inflammatory index

FFQ-derived dietary data were used to calculate DII scores for each study participant. A complete description of the DII is available elsewhere⁽²⁰⁾. Briefly, on the basis of a search of the literature from 1950 to the end of 2010, we identified forty-five food parameters among foods, nutrients and other food components that were associated with six plasma inflammatory markers (IL-1 β , IL-4, IL-6, IL-10, TNF- α and CRP). We defined a specific DII score for each food parameter on the basis of the literature review and taking into account the quality and number of published papers (1943 articles were reviewed and scored).

For each study participant, the dietary data were first linked to a global database that was developed on the basis of eleven data sets from around the world and thus provides a robust estimate of the means and standard deviations of these fortyfive parameters⁽²⁰⁾. Each participant's exposure relative to the 'standard global mean' was expressed as a Z score that was derived by subtracting the 'standard global mean' from the amount reported and then dividing this value by its sp. To minimise the effect of 'right skewing', this value was then converted to a centred percentile score. The participant's DII score was computed by multiplying this value by the specific DII score for each food parameter and by summing together all forty-five values according to the following formula: $DII = b_1 \times n_1 + b_2 \times n_2 + \dots + b_{45} \times n_{45}$, where b_i refers to the literature-derived inflammatory effect score for each of the evaluated food parameter and n_i refers to the food parameterspecific centred percentile, which were derived from the dietary data, per each *i* from 1 to 45. A higher DII score indicates a more pro-inflammatory diet. The DII computed on this study's FFQ includes data on thirty-one of the forty-five food parameters comprising the DII, including carbohydrate, protein, fat, alcohol, fibre, cholesterol, SFA, MUFA, PUFA, n-3, n-6, niacin, thiamin, riboflavin, vitamin B₆, Fe, Zn, vitamin A, vitamin C, vitamin D, vitamin E, folic acid, β -carotene, anthocyanidins, flavanol, flavonol, flavonones, flavones, isoflavones, caffeine and tea. A flow chart of the DII methodology is shown in Fig. 1.

325

MS British Journal of Nutrition

N. Shivappa et al.

Review of articles published from 1950 to 2010 resulting in 1943 studies linking a total of 45 food parameters with inflammatory biomarkers.



A score for each food parameter was calculated giving:

+1 to each article if the effects were pro-inflammatory (significantly increased IL-1 β , IL-6, TNF- α or CRP, or decreased IL-4 or IL-10),

-1 if the effects were anti-inflammatory (significantly decreased IL-1 β , IL-6, TNF- α or CRP, or increased IL-4 or IL-10).

0 if the food parameter did not produce any significant change in the inflammatory marker.



The score for each food parameter was weighted according to the study design. The weights were 10 (experimental design), 8 (observational), 7 (case-control), 6 (cross-sectional), 5 (experimental with animals), 3 (cell culture).



A food parameter-specific overall inflammatory effect score was calculated by substracting the anti-inflammatory fraction from the pro-inflammatory fraction. This score was corrected if the total weighted number of articles was <236. In these cases the raw overall inflammatory score is multiplied by the total weighted number of articles divided by 236.



Z-score and centred-percentiles for each of the 31 food parameters for each participant of this study were calculated based on the average and standard deviation for each food parameter obtained from the global database which was created from the consumption of the original 45 food parameters fron 11 countries from around the world.



The centred percentile for each food parameter was multiplied by the the respective 'overall food parameter-specific inflammatory effect score' to obtain the 'food parameter-specific DII score'.



All of the 'food parameter-specific DII scores' are summed to create the 'overall DII score' for each individual.

Fig. 1. Sequence of steps in creating the dietary inflammatory index (DII) in the Italian hepatocellular cancer case-control study. CRP, C-reactive protein.

The DII was analysed both as a continuous variable (i.e. for a one-unit increment in the DII corresponds to approximately 7% of its global range) and by tertiles of exposure, determined on

the basis of the entire study population. The DII also was examined across the strata of selected factors such as age, education, BMI and physical activity using the ANOVA test for

327

continuous variables or the χ^2 test for categorical variables. OR and the corresponding 95% CI were estimated using logistic regression models, adjusting only for age and then additionally for serum markers of hepatitis B and C infection, sex, study centre, education (<7, 7–11 and >12 years), BMI (<25.0, 25.0–29.9 and $\geq 30.0 \text{ kg/m}^2$) and tobacco smoking (never smokers, ever smokers <15 cigarettes/d and ever smokers \geq 15 cigarettes/d). Energy adjustment for DII was performed using the residual method⁽³⁷⁾. Alcohol was not added as a covariate because alcohol is one of the components used for DII calculation. Linear tests for trend were performed using the median value within each tertile as an ordinal variable. The test for heterogeneity was carried out by including the interaction terms between DII and hepatitis status and sex in the model. Stratified analyses were carried out by hepatitis B and C infection status and sex. Statistical analyses were performed using SAS[®] 9.3 (SAS Institute Inc.).

Results

Controls were more often females and were younger compared with cases. Cases were more likely than controls to have a low level of education, to smoke cigarettes and to report heavy alcohol drinking⁽³⁸⁾. The mean DII value among cases was 0.24 (sp 1.40), and among controls it was -0.11 (sp 1.37), indicating a slightly more pro-inflammatory diet for cases. Control characteristics across tertiles of DII are provided in Table 1. There were some differences in socio-demographic, anthropometric and lifestyle habits across DII tertiles. Compared with controls in tertile 1, participants in the highest tertile were younger and more likely to be former smokers, to have a higher BMI and to have educational attainment of <7 years.

OR and 95% CI of HCC according to tertiles of DII and continuous DII are shown in Table 2. When analyses were carried out using continuous DII, a significant positive association was observed with HCC risk (multivariable OR 1.30; 95% CI 1.05, 1.60). When fit as tertiles, participants in tertile 3 were at higher risk for having HCC compared with participants in tertile 1 (OR_{tertile 3 v. 1} 2.43; 95% CI 1.27, 4.68; $P_{\text{trend}} = 0.03$).

The interaction term for DII-by-sex was significant (*P* value = 0.04) but not for hepatitis infection (*P*=0.09). Still, we carried out a stratified analyses. When stratified by sex, strong associations were observed only among males (OR_{tertile 3 v. 1} 3.60; 95% CI 1.65, 7.87; *P*_{trend} = 0.001). When stratified by hepatitis infection status, the DII was strongly associated with HCC among hepatitis B surface antigen (HBsAg)- and anti-hepatitis virus C2 (HCV2)-negative participants (i.e. the hepatitis-negative group), *n* 38 (OR_{tertile 3 v. 1} 4.18; 95% CI 1.53, 11.39; *P*_{trend} = 0.02). No statistically significant associations were observed among HBsAg + and/or anti-HCV + group, *n* 147 (Table 3).

Discussion

In this case–control study, consuming a more pro-inflammatory diet – as reflected in higher DII scores – was associated with an increased risk for HCC. Previous results obtained in this case–control study showed deleterious effects of higher dietary

Table 1. Participants' characteristics across tertiles of dietary inflammatory index among 404 controls, Italy, 1999–2002(Percentages; mean values and standard deviations)

Characteristics	Tertile 1 < - 0.67	Tertile 2 -0.67-0.62	Tertile 3 >0.62	<i>P</i> *
Age (years) (% in each category)				0.14
<55	15.4	20.6	26.9	
55–64	31.5	32.3	20.2	
65–74	38.3	30.9	37.8	
≥75	14.8	16.2	15.1	
Sex (% in each category)				0.83
Males	67.1	69.1	70.6	
Females	32.9	30.9	29.4	
BMI (kg/m ²)				0.30
Mean	26.6	27.0	27.2	
SD	4.3	3.7	4.0	
Smoking (% in each category)		-		0.01
Non-smoker	33.6	32.8	33-1	
Former smoker	47.0	42.5	31.4	
Current cigarette smokers of 1-14 CPD	12.1	14.2	14-1	
Current cigarette smokers of >15 CPD	7.4	10.5	21.5	
Education (vears) (% in each category)				0.22
<7	49.7	56.6	62.3	
7–11	24.8	21.3	22.7	
>11	25.5	22.1	15.1	
Hepatitis infection (% in each category)†				0.70
No	90.6	87.5	89-1	
Yes	9.4	12.5	10.9	
Alcohol consumption (% in each category)	0.1	0		0.81
Abstainer	14.1	16.2	16.0	
Current	77.9	72.0	73.9	
Former	8.0	11.8	10.1	

CPD, cigarettes per d.

* ANOVA was used for continuous variables and χ^2 test was used for categorical variables.

† Hepatitis was defined as hepatitis B surface antigen and/or anti-hepatitis C virus positivity.

Table 2. Risk for hepatocellular cancer and dietary inflammatory index (DII) both as continuous and expressed as tertiles, among 185 cases and 404 controls, Italy, 1999–2002

(Odds ratios and 95% confidence intervals)

	DII								
	Tertile 1 <-0.67 Tertile 2		-0.67-0.62	Tertile	Tertile 3 >0.62		DII (continuous)*		
	OR	OR	95 % CI	OR	95 % CI	Р	OR	95 % CI	Р
Cases/controls Model 1† Model 2§	47/149 1‡ 1‡	60/134 1.63 1.21	1.02, 2.58 0.63, 2.31	78/121 2·67 2·43	1.67, 4.27 1.27, 4.68	<0.0001 0.03	185/404 1·29 1·24	1·12, 1·49 1·02, 1·51	0.0003 0.03

HBsAg, hepatitis B surface antigen; HVC, hepatitis virus C.

* One unit increase corresponding to approximately 12 % of its range in the current study.

† Estimated using age- and energy-adjusted logistic regression model. Energy intake (kcal), energy from alcohol excluded)

‡ Reference.

§ Additionally adjusted for age, sex, centre, BMI (<25·0, 25·0–29·9 and ≥30·0 kg/m²), smoking (non-smoker, current smoker, former smoker), education (<7, 7–11 and ≥12 years), hepatitis viruses (HBsAg+ and/or anti-HCV+ ν. HBsAg- and anti-HCV-).</p>

Table 3. Risk for hepatocellular cancer (HCC) and dietary inflammatory index (DII) both as continuous and expressed as tertiles, in separate strata of hepatitis C and/or B virus infection and sex (185 cases and 404 controls), Italy, 1999–2002 (Odds ratios and 95 % confidence intervals)

	DII								
	Tertile 1 <-0.67*	Tertile 2 -0.67-0.62		Tertile 3 >0 62			DII (continuous)†		
HCC by hepatitis viral status	OR	OR	95 % CI	OR	95 % CI	Р	OR	95 % Cl	Р
HBsAg- and anti-HCV-negative (cases/controls = 38/360)	9/135	9/117		20/108					
Model 1‡	1.00	1.24	0.43, 3.58	4.18	1.53, 11.39	0.02	1.64	1.19, 2.24	0.002
HBsAg + and/or anti-HCV-positive $(cases/controls = 147/44)$	38/14	52/17		57/13					
Model 1‡	1.00	0.98	0.35, 2.77	1.76	0.57, 5.48	0.59	1.07	0.77, 1.49	0.68
Sex									
Male (cases/controls = 148/278)	34/100	49/93		65/85					
Model 1‡	1.00	1.56	0.73, 3.30	3.60	1.65, 7.87	0.001	1.39	1.10, 1.76	0.006
Female (cases/controls = 36/126)	13/49	11/41		12/36					
Model 1‡	1.00	0.40	0.08, 2.10	0.44	0.07, 2.80	0.37	0.86	0.53, 1.40	0.54

HBsAg, hepatitis B surface antigen; HVC, hepatitis virus C.

* Reference.

NS British Journal of Nutrition

† One unit increase corresponding to approximately 12 % of its range in the current study.

‡ Estimated using logistic regression model adjusting for energy intake (kcal, energy from alcohol excluded), sex, centre, BMI (<25.0, 25.0–29.9 and ≥30.0 kg/m²), smoking (nonsmoker, current smoker, former smoker), education (<7, 7–11 and ≥12 years).
</p>

glycaemic load (GL); the GL is a ranking of specific carbohydrate-rich foods based on the postprandial blood glucose response, which is expressed as a percentage of the glycaemic response to an equivalent amount of available carbohydrates from a reference food (e.g. white bread or glucose)⁽³⁹⁾. From our prior research, we know that carbohydrate has a strong pro-inflammatory DII score⁽²⁰⁾. A diet rich in linoleic acid-containing foods (e.g. white meats and fish) and β -carotene was inversely related to HCC risk⁽⁴⁰⁾. There is a favourable effect of coffee, although neither decaffeinated coffee nor tea^(19,41), on the risk for HCC. All of these nutrients and food items have anti-inflammatory values in the DII scoring⁽²⁰⁾.

Previous studies that have examined the effect of specific food items on HCC have reported an increased risk for high consumption of red meat, SFA⁽⁴²⁾ and excessive alcohol drinking^(9,43,44), whereas no association was observed with SFA in a European cohort⁽⁴⁵⁾ and with red meat in this case–control study⁽⁴⁶⁾. There is strong evidence showing coffee to be protective

against $\text{HCC}^{(47-50)}$. Inverse associations also have been observed for white meat and fish^(42,51). A report from two cohort studies in China showed vitamin E to be protective against $\text{HCC}^{(18)}$.

Two studies have been conducted to examine various other dietary patterns and indices in relation to HCC^(16,17). In the National Institute of Health-American Association of Retired Professionals cohort, after adjustment for multiple confounders, significant inverse associations were observed between HCC incidence and the Healthy Eating Index-2005 and the alternate Mediterranean Diet Score⁽¹⁶⁾. In the Shanghai Women's and Men's Health Studies, fruit- and meat-based dietary patterns were not associated with liver cancer risk; however, a vegetable-based dietary pattern was associated with reduced liver cancer risk⁽¹⁷⁾. In a large European cohort only vegetable consumption was observed with fruit intake⁽⁵²⁾.

One of the possible mechanisms for the observed positive association between the DII and HCC is through the indirect effect of pro-inflammatory diet due to enhanced production of

329

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the tumour-promoting cytokines IL-6 and TNF, which cause hepatic inflammation and activation of the oncogenic transcription factor, signal transducer and activator of transcription 3 (STAT3)⁽⁵³⁾. Consumption of a diet rich in SFA increases pro-inflammatory responses through the process of peroxidation of lipids in cells, which then exacerbates liver injury leading to HCC⁽⁵⁴⁾. Our finding that the positive association with the DII was restricted to the HBsAg- and anti-HCV2-negative group may indicate either that these viral infections operate through mechanisms independent of inflammation or that other factors related to these infections overwhelm inflammatory effects. Consequently, it may be that the effect of DII is more apparent in the absence of such infections. In the face of intense virus-induced inflammation, pro-inflammatory diet might have little impact. As this is the first study showing this result, the association has to be explored further in other studies.

Because of the relative rarity of the disease, the study of HCC in high-income countries has traditionally only been feasible using case-control designs. However, pooling of results across cohort studies make it feasible to explore HCC in combined data of several cohort studies too^(50,52). In this case-control study, both cases and controls came from comparable catchment areas and were interviewed by uniformly trained interviewers in their respective hospital settings. To limit possible sources of bias, we included in the control group patients admitted for a wide spectrum of acute, non-neoplastic conditions, unrelated to the major risk factors for HCC. The practically complete participation rate and the comparable catchment areas of cases and controls contributed to reduce any potential selection bias. Cases may recall history of other diseases more frequently than did controls. However, the hospital setting should have improved the comparability of information, as cases and controls are interviewed under similar conditions. Participants were unaware of any particular studyrelated hypothesis in relation to diet and HCC, thereby reducing potential selection and information bias⁽⁵⁵⁾. The FFQ was satisfactorily reliable⁽³⁴⁾ and validated⁽³⁵⁾.

Other limitations are the non-availability of the remaining thirteen food parameters for the DII calculation. DII calculated from these thirty-one food parameters has not been validated with inflammatory markers. However, in previous validation studies, the DII has been calculated from food parameters ranging from 17 to 44 and produced reasonably uniform results with minimal reduction in predictive ability^(21,22). Moreover, there could be a possible overestimation due to the inclusion of food items in the DII calculation that are also a source of nutrients. It also should be noted, however, that each of these food items has an inflammatory effect score, which is derived from an extensive review of the literature looking at the association between these foods and inflammation. Another limitation could be the absence of evidence of DII being associated with inflammatory markers in this study, although the case-control nature of the study is not amenable to adding such measures during the aetiologically relevant period in any event. Our results were adjusted for the majority of known risk factors for HCC, including education, tobacco smoking, BMI and energy intake.

In conclusion, Italian men and women who consumed a more pro-inflammatory diet high in components of dietary GL and low in PUFA, vitamin A and β -carotene were at increased risk for HCC compared with those who consumed a more anti-inflammatory diet. The results suggest that encouraging intake of more anti-inflammatory dietary factors – such as coffee, white meat in place of red meat and plant-based foods rich in vitamin E and phytochemicals – and reducing intake of pro-inflammatory factors – such as red meat rich in SFA – may be a strategy for reducing risk for HCC.

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A. Z., M. M., C. L. V., S. F. and D. S. designed and conducted the case–control study; N. S. conducted the analyses and wrote the first draft of the manuscript; M. R., J. R. H., C. L. V., J. P., A. C., M. M., S. F., and A. Z. provided suggestions and revised the manuscript. All authors approved the final version of the manuscript.

The authors declare that there are no conflicts of interest.

J. R. H. owns controlling interest in Connecting Health Innovations LLC (CHI), a company planning to license the right to his invention of the dietary inflammatory index from the University of South Carolina in order to develop computer and smart phone applications for patient counselling and dietary intervention in clinical settings. Dr N. S. is an employee of CHI.

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N. Shivappa et al.

330

NS British Journal of Nutrition

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331